

The role, or non-role, of constraints in the forecasting of mortality

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The purpose of this paper is to show

- there is **no identifiability problem**
- **forecasting does not depend** on the choice of constraints

Mortality data

		Year			
		1	2	3	n_y
Age	1	n_x			$n_x + n_y - 1$
	2				$= n_c$
	3				
	n_x	3	2 3	1 2 3	

Deaths : D
Exposures : E
 $D, E : n_x \times n_y$

Age of death = rows, Year of death = columns, Year of birth = diagonals

Illustrative Data

ONS: UK males: Ages: 50-104; Years: 1971-2015.

$$n_x = 55; n_y = 45; n_c = 99; N = n_x n_y = 2475.$$

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where $c(i,j) = n_x - i + j$.

- The **Age-Period-Cohort-Improvement (APCI)** model is

$$\log \mu_{i,j} = \alpha_i + \kappa_j + \gamma_{c(i,j)} + \beta_i(y_j - \bar{y}).$$

This is used by the CMI to parameterise its forecasting spreadsheet.

Generalized linear models (GLMs)

Let $\mathbf{D} = (d_{i,j})$ and $\mathbf{E} = (e_{i,j})$ and assume

$$d_{i,j} \sim \mathcal{P}(e_{i,j}\mu_{i,j})$$

where, for example, in AP model $\log \mu_{i,j} = \alpha_i + \kappa_j$.

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In vector form

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The AP, APC and APCI are all GLMs (overdispersion is ignored).

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Model matrix \mathbf{X} is $N \times (n_x + n_y)$ and has rank $n_x + n_y - 1$

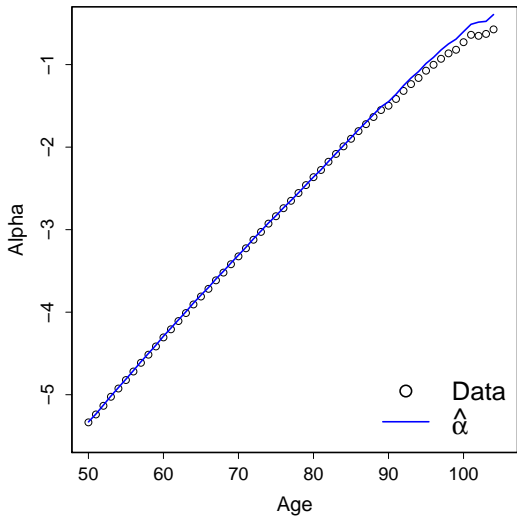
\Rightarrow parameters are not uniquely estimable.

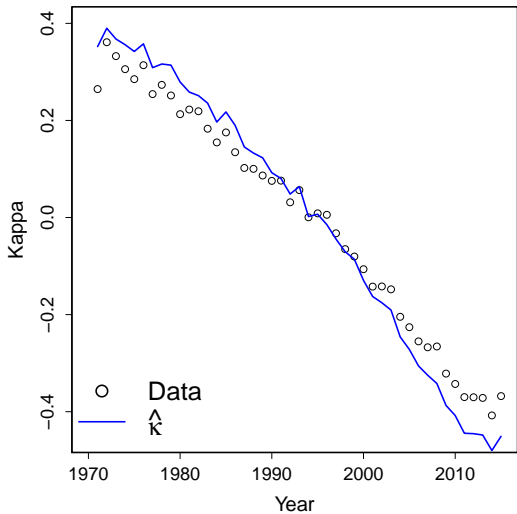
Identifiability problem

Constraints in AP model

Standard constraint: $\sum \kappa_j = 0$

⇒ parameters are uniquely estimable.





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BUT

Random constraints in AP model

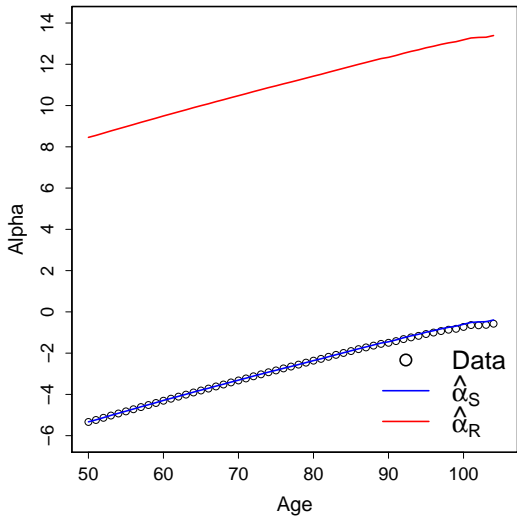
Let $\theta' = (\alpha', \kappa)'$.

$$\text{Random constraint: } \sum_1^{n_x+n_y} u_i \theta_i = 0$$

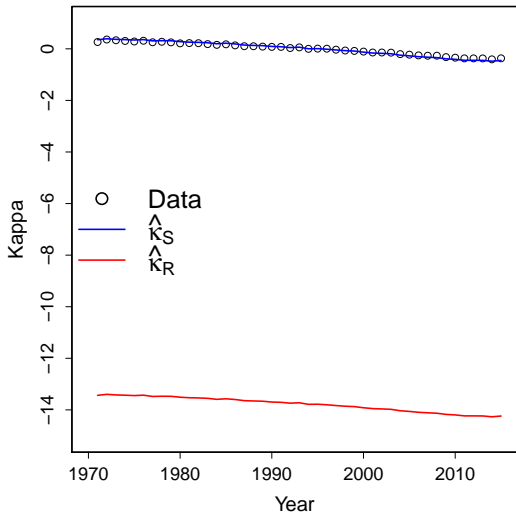
where $U_i \sim \mathcal{U}(0, 1)$.

\Rightarrow parameters are uniquely estimable.

Standard: $\hat{\alpha}_S$, Random: $\hat{\alpha}_R$



Standard: $\hat{\kappa}_S$, Random: $\hat{\kappa}_R$



Properties

Standard (centred) estimates: $\hat{\theta}_S = (\hat{\alpha}'_S, \hat{\kappa}'_S)'$

Random estimates: $\hat{\theta}_R = (\hat{\alpha}'_R, \hat{\kappa}'_R)'$.

Define

$$\Delta \hat{\alpha} = \hat{\alpha}_S - \hat{\alpha}_R, \quad \Delta \hat{\kappa} = \hat{\kappa}_S - \hat{\kappa}_R$$

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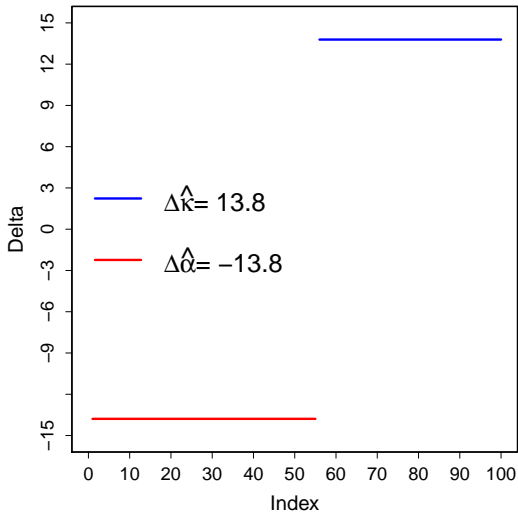
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- Here $k = -13.8$.

$$\Delta \hat{\alpha} = \hat{\alpha}_S - \hat{\alpha}_R, \text{ etc}$$



Forecasting

Forecasting with ARIMA model, e.g., random walk with drift, is **invariant wrt choice of constraints**.

Age-Period-Cohort Model

The APC model is

$$\log \mu_{i,j} = \alpha_i + \kappa_j + \gamma_{c(i,j)}$$

where $c(i,j) = n_x - i + j$.

Constraints

Standard (Cairns et al, 2009):

$$\sum \kappa_j = \sum \gamma_c = \sum w_c \gamma_c = 0$$

where w_c is the cohort index, $w_c = 1, \dots, n_c$.

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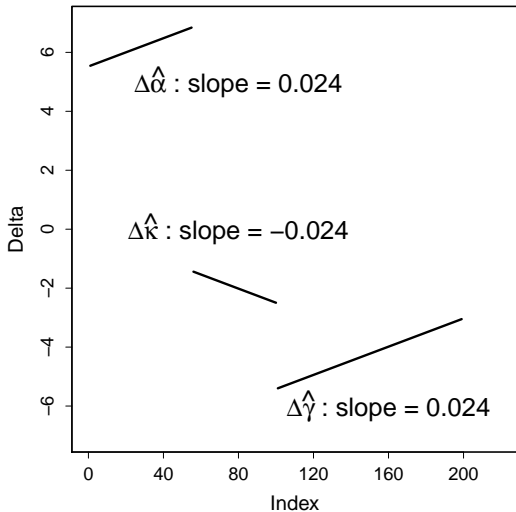
where w_c is the cohort index, $w_c = 1, \dots, n_c$.

Random: Let $\theta = (\alpha', \kappa', \gamma')'$.

$$\sum u_{1,j} \theta_j = \sum u_{2,j} \theta_j = \sum u_{3,j} \theta_j = 0$$

where the $u_{i,j}$, $i = 1, 2, 3$, $j = 1, \dots, n_x + n_y + n_c$, are $\mathcal{U}(0, 1)$.

$$\Delta \hat{\alpha} = \hat{\alpha}_S - \hat{\alpha}_R, \text{ etc}$$



Forecasting

- Forecasting with ARIMA model, e.g., random walk with drift, is **invariant wrt choice of constraints**.

Age-Period-Cohort-Improvement (APCI) Model

The model is

$$\log \mu_{i,j} = \alpha_i + \kappa_j + \gamma_{c(i,j)} + \beta_i(y_j - \bar{y})$$

and forms the basis for the CMI's current forecasting spreadsheet.

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Model matrix \mathbf{X} is $N \times (3n_x + 2n_y - 1)$ and rank $3n_x + 2n_y - 6$ and five (5) constraints are required to bring about identifiability.

Constraints

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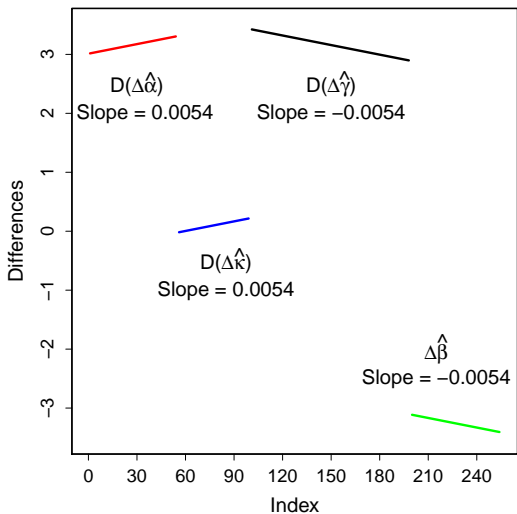
Random: Let $\theta = (\alpha', \kappa', \gamma', \beta')'$.

$$\sum u_{i,j} \theta_j = 0, \quad i = 1, \dots, 5, \quad j = 1, \dots, 2n_x + n_y + n_c,$$

where the $u_{i,j}$, are $\mathcal{U}(0, 1)$.

$$\Delta \hat{\alpha} = \hat{\alpha}_S - \hat{\alpha}_R, \text{ etc}$$

D = difference operator



Forecasting

- Forecasting with ARIMA model is invariant wrt choice of constraints provided $d \geq 3$ in ARIMA model.

Smoothing

In AP and APC models smooth α . Set

$$\alpha = B_a a$$

where B_a is a regression matrix of B -splines for age.

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In APCI model additionally smooth β . Set

$$\beta = B_a b$$

Use method of P -splines (Eilers & Marx, 1996).

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- Fitted and forecast values are invariant wrt choice of constraints.
- Order of
 - penalty for smoothing and
 - differencing in ARIMA modelmust be sufficiently large (see Currie (in preparation) for details).

References

1. Cairns, Blake, Dowd et al (2009) A quantitative comparison of stochastic mortality models using data from England and Wales and the United States. *North American Actuarial Journal*, **13**, 1–35.
2. Currie (2013) Smoothing constrained generalized linear models with an application to the Lee-Carter model. *Statistical Modelling*, **13**, 69–93.
3. Currie (in preparation) Constraints, the identifiability problem and the forecasting of mortality.
4. Richards, Currie, Kleinow & Ritchie (to appear). A stochastic implementation of the APCI model for mortality projections.