

Non-nested model selection
in unstable environments

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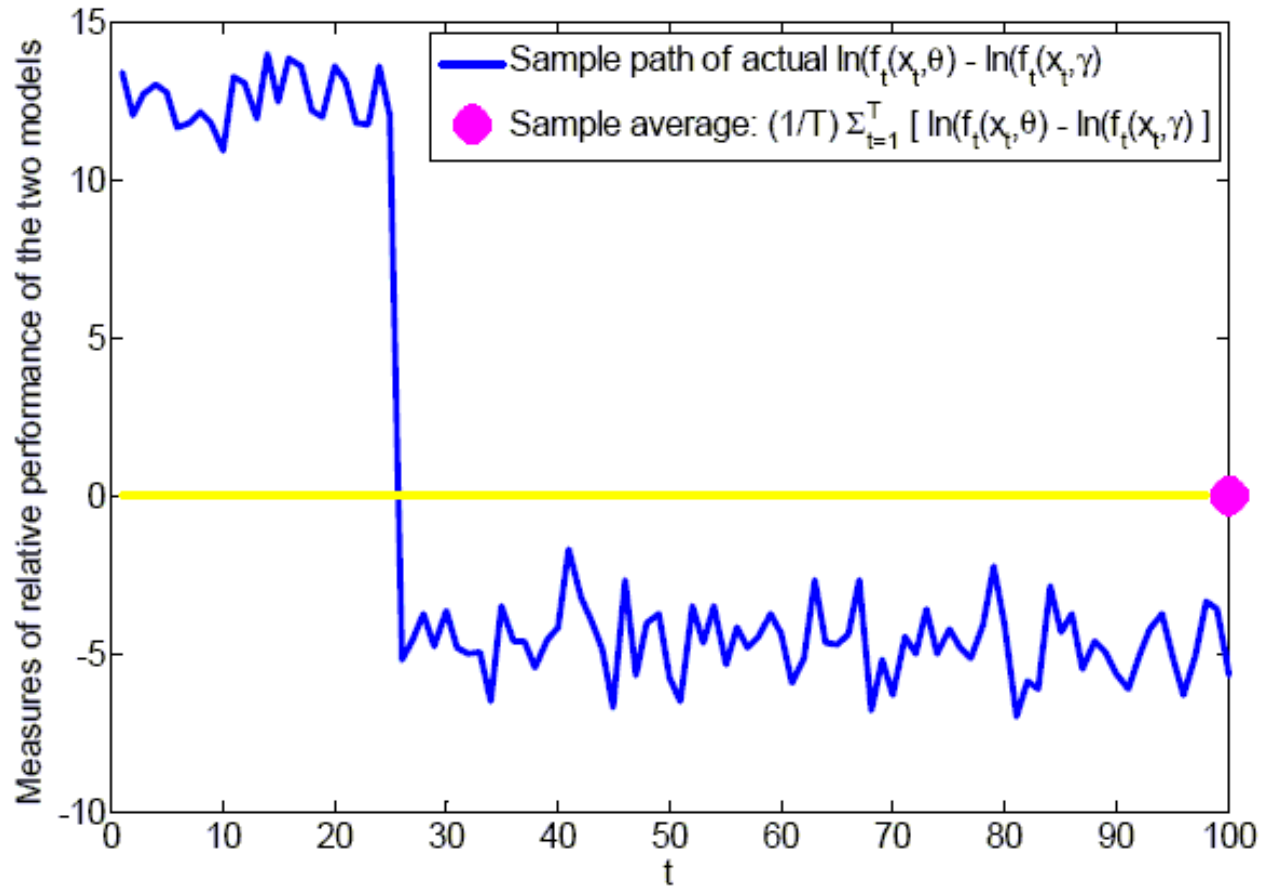
Motivation

- The problem: select between two competing models, based on how well they fit the data
- Both models are possibly misspecified
- Widely documented instability in macro and financial time series \implies the relative performance of the models may also be unstable
- Existing techniques for model selection compare the *average* performance of the models over historical samples \implies loss of information if relative performance varies over time
- Our goal: propose formal techniques to test whether the relative performance of two models is stable

Example

- Model 1 is better than model 2 (= higher likelihood) at the beginning of the sample and then the order switches.
- One would like to choose model 2 at the end of the sample, but a measure of average performance may find that the models are equivalent.

Example (cont.)



Contributions

- We propose new model selection tests to evaluate whether the relative performance of two competing models varies over time. Two tests:
 - **Fluctuation test** to analyze the evolution of the models' relative performance over **historical** samples
 - **Sequential test** to monitor the models' relative performance in **real time**, as new data becomes available
- Valid under general conditions: data instability; models can be nonlinear, multivariate and misspecified; ML estimation
- Empirical application: evaluate time-variation in performance of Smets and Wouters' (2003) DSGE model relative to BVARs.

Outline of the talk

- Fluctuation test
- Sequential test
- Monte Carlo evidence
- Empirical application to DSGE vs. BVAR
- Conclusion

Related literature

Fluctuation test:

- Vuong, 1989 \implies test for equal full-sample *average* relative performance of non-nested, misspecified models (iid data)
- Rivers and Vuong, 2002 \implies heterogeneous + dependent data but tests that models are asymptotically equivalent
- Diebold- Mariano, 1995; West, 1996; McCracken 2000 etc. \implies test for equal out-of-sample *average* relative performance of non-nested, misspecified models
- Giacomini and White, 2006 \implies test if out-of-sample relative performance is related to explanatory variables

Related literature (cont.)

- Rossi, 2005 \implies test if parameters are stable + significant (\sim nested model selection under instability, but assumes correct specification)
 - We test a different H_0 : the relative performance is *equal at each point in time* + consider non-nested, possibly misspecified models
- Tests for parameter instability (Brown, Durbin and Evans, 1975, Andrews, 1993, Bai, 1998 etc. etc.)
 - We borrow from this literature, but focus on stability of relative performance rather than model's parameters (relative performance may be stable even though parameters change)

Related literature

Sequential test:

- Chu, Stinchcombe and White (1996) \implies real-time *parameter* stability within one model
- Inoue and Rossi (2005) \implies assess if variable has real-time predictive content for another variable
 - We ask a different question (compare non-nested models based on measures of fit) and have more general assumptions (e.g., both models misspecified)

Fluctuation test - Set-up

- Select a model for vector y_t using other variables z_t . Let $x_t = (y_t', z_t')'$. Historical sample (x_1, \dots, x_T) .
- Two competing models, estimated by ML.
- Idea: re-estimate models recursively starting from observation $R < T$ using either an expanding sample ("recursive scheme") or a rolling samples of size R ("rolling scheme")
- At each time t , measure relative performance as $Q_t(\hat{\theta}_t) - Q_t(\hat{\gamma}_t)$ where $Q_t(\hat{\theta}_t)$ is the average log-likelihood over the estimation sample

Fluctuation test - Rolling scheme

Performance measure for model 1: $Q_t(\hat{\theta}_t) = R^{-1} \sum_{j=t-R+1}^t \ln f(x_j; \hat{\theta}_t)$

$$\begin{array}{l}
 \begin{array}{c} \text{-----} \\ 1 \qquad \qquad R \end{array} \rightarrow Q_R(\hat{\theta}_R) - Q_R(\hat{\gamma}_R) \\
 \begin{array}{c} \text{-----} \\ 2 \qquad \qquad R+1 \end{array} \rightarrow Q_{R+1}(\hat{\theta}_{R+1}) - Q_{R+1}(\hat{\gamma}_{R+1}) \\
 \begin{array}{c} \text{-----} \\ T-R+1 \qquad T \end{array} \rightarrow Q_T(\hat{\theta}_T) - Q_T(\hat{\gamma}_T)
 \end{array}$$

Fluctuation test - Null hypothesis

- Null hypothesis: $H_0 : EQ_t(\theta_t^*) - EQ_t(\gamma_t^*) = 0$ for all $t = 1, \dots, T$,
- θ_t^* and γ_t^* are the pseudo-true parameters. E.g., for the recursive scheme,

$$\theta_t^* = \arg \max_{\theta} E \left(t^{-1} \sum_{j=1}^t \ln f(x_j; \theta) \right)$$

- Note that the parameters of the models may be unstable under the null hypothesis

Fluctuation test - Implementation

- Compute sequences of statistics for $t = R, \dots, T$

$$F_t^{rec} = \hat{\sigma}_t^{-1} \sqrt{t} \left(Q_t(\hat{\theta}_t) - Q_t(\hat{\gamma}_t) \right)$$
$$F_t^{roll} = \hat{\sigma}_t^{-1} \sqrt{R} \left(Q_t(\hat{\theta}_t) - Q_t(\hat{\gamma}_t) \right),$$

$\hat{\sigma}_t^2$ is a HAC estimator of the asymptotic variance $\sigma_t^2 = \text{var}(\sqrt{t} (Q_t(\theta_t^*) - Q_t(\gamma_t^*)))$

- Does the sample path of F_t significantly deviate from its hypothesized value of 0?

Fluctuation test - Recursive scheme

- Intuition: if the models are non-nested, for a particular t , F_t^{rec} can be approximated by a $N(0, 1)$ under $H_0 \implies$ the sample path of F_t^{rec} behaves like a Brownian motion that starts at zero at time $t = R$.

\implies can derive boundary lines that are crossed with given probability α . For each $t = R, \dots, T$ and significance level α :

$$c_{\alpha,t}^{rec} = \pm k_{\alpha}^{rec} \sqrt{\frac{T-R}{t}} \left(1 + 2 \frac{t-R}{T-R} \right)$$

(Same as for the CUSUM test of Brown, Durbin and Evans, 1975)

- Typical values of $(\alpha, k_{\alpha}^{rec})$ are $(0.01, 1.143)$, $(0.05, 0.948)$ and $(0.10, 0.850)$

Fluctuation test - Rolling scheme

- For the rolling scheme, the sample path of F_t^{roll} behaves like the increments of a Brownian bridge \implies can derive boundary lines that are crossed with given probability α by simulation

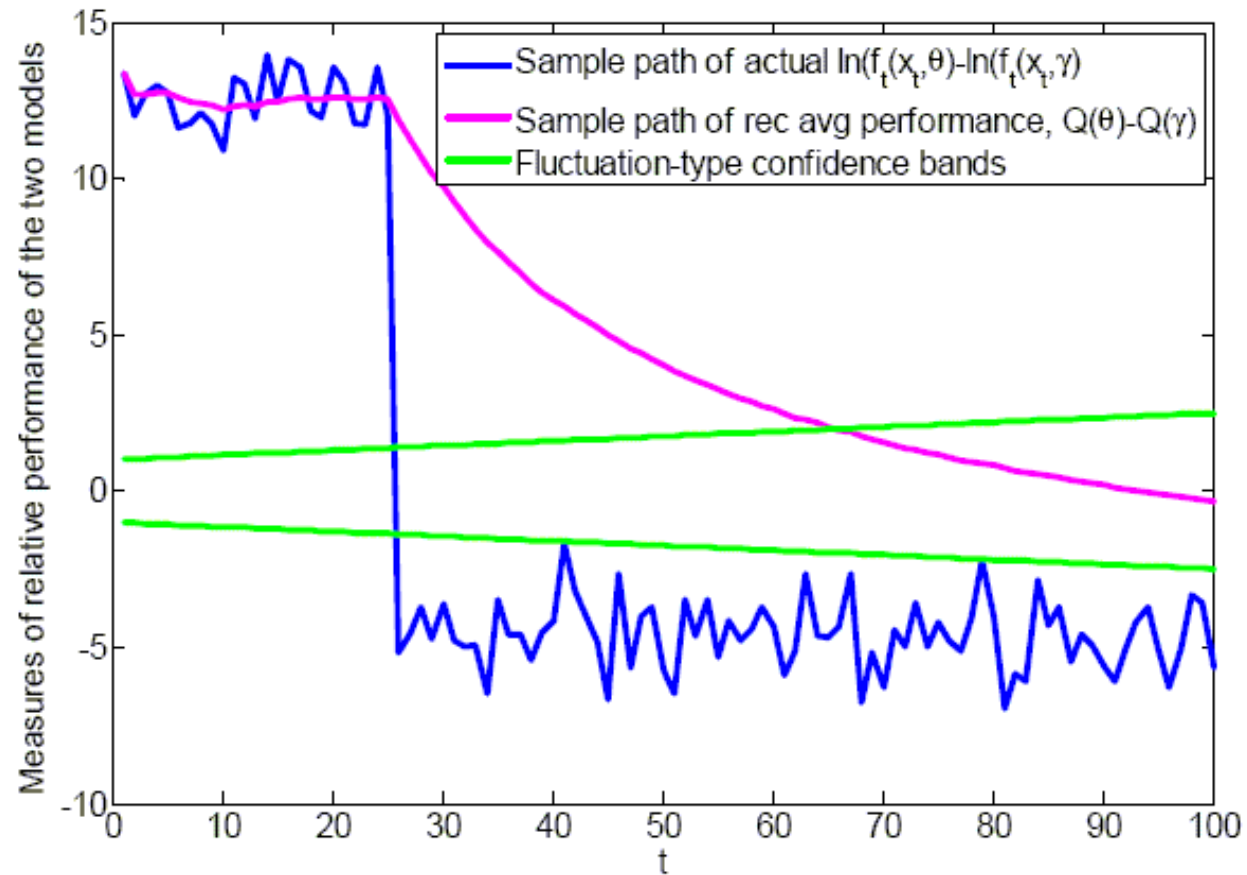
$$c_{\alpha,t}^{roll} = \pm k_{\alpha}^{roll}$$

- k_{α}^{roll} depends on R/T . We give a table with k_{α}^{roll} for typical values of a and R/T

Fluctuation test - Assumptions

- Standard assumptions that guarantee that a FCLT holds for $T^{-1/2} \sum_{j=1}^t \ln f(x_j; \theta)$
- Parameter and data instability allowed under H_0
- Data have short memory (\implies no unit roots)
- σ_t^2 is not $o(1)$ (rules out nested models)
- $R/T \rightarrow \rho$ finite and positive

Fluctuation test in practice (recursive)



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Sequential test

- Monitor the model-selection decision in the post-historical sample period
- Suppose model 1 was best over the historical sample up to time T :

$$EQ_T(\theta_T^*) - EQ_T(\gamma_T^*) > 0$$

- Test hypothesis that model 1 continues to be the best for all subsequent periods:

$$H_0 : EQ_t(\theta_t^*) - EQ_t(\gamma_t^*) \geq 0 \text{ for } t = T + 1, T + 2, \dots,$$

against $H_1 : EQ_t(\theta_t^*) - EQ_t(\gamma_t^*) < 0$ at some $t \geq T$.

- Doing a sequence of full-sample Vuong's (1989) tests rejects too often \implies find critical values that control the overall size of the procedure

Sequential test

- Compute sequence of test statistics for $t = T + 1, T + 2, \dots$

$$J_t = \hat{\sigma}_t^{-1} \sqrt{t} \left(Q_t(\hat{\theta}_t) - Q_t(\hat{\gamma}_t) \right)$$

- As in Chu, Stinchcombe and White (1995), the critical value at time t for a level α test is:

$$c_{\alpha,t} = -\sqrt{k_\alpha + \ln(t/T)}$$

- Typical values of (α, k_α) are $(0.05, 2.7955)$ and $(0.10, 2.5003)$.

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Monte Carlo evidence

- Compare our test to Vuong's (1989) in-sample test and to West's (1996) out-of-sample test

- DGP with parameter variation:

$$y_t = \beta_t x_t + \gamma_t z_t + \varepsilon_t, \quad t = 1, 2, \dots, T, \quad T = 400$$

$$\beta_t = 1 + \beta \cdot \mathbf{1}(200 < t \leq 250) + (1 - \beta) \cdot \mathbf{1}(t > 250)$$

$$\gamma_t = 1 + \gamma \cdot \mathbf{1}(200 < t \leq 250) + (1 - \gamma) \cdot \mathbf{1}(t > 250).$$

- Model 1: $y_t = \beta x_t + u_{1t}$. Model 2: $y_t = \gamma z_t + u_{2t}$.
- Size: $\beta = \gamma = 0.5 \implies$ models are equally good.
- Power: $\beta = 0.95, \gamma = 0.4 \implies$ time variation in relative performance

Monte Carlo evidence

Table 2. Empirical rejection frequencies of nominal 5% tests.

		F_t^{rec}	Vuong	West
(a) Historical sample	Size	0.051	0.047	0.044
	Power	0.449	0.047	0.026
(b) Post-historical sample	t/T	J_t	Vuong	
	1.5	0.010	0.121	
	1.75	0.020	0.152	
	2	0.032	0.179	

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Application: SW's DSGE vs. BVAR

- Smets and Wouters (2003) (SW): “An estimated DSGE model of the Euro Area”: estimation of a 7-equation linearized DSGE model with sticky prices and wages, habit formation, capital adjustment costs and variable capacity utilization.
- Their findings:
 - the DSGE model has comparable fit to that of atheoretical VARs
 - the parameter estimates have the expected sign

Open questions

- Have the parameters been stable? Perhaps not. Possible *structural changes* in the economy (European union introduction, productivity changes, etc.)
- If the parameters have changed \implies the performance of the DSGE model may have changed too... so SW's result only holds on average
- Can we say that the DSGE's performance has been *stable over time*, and *significantly better* than that of the VAR?

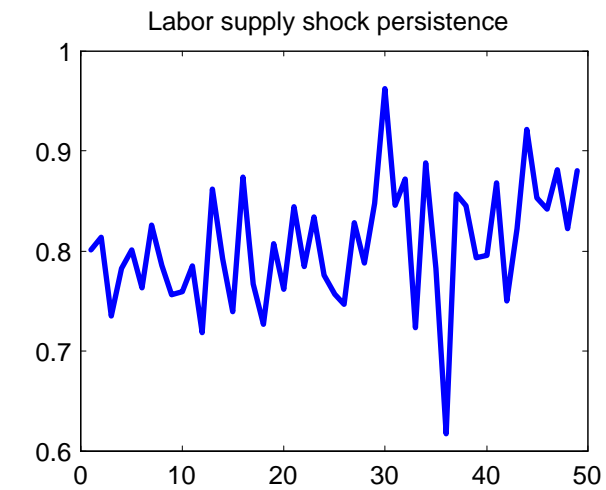
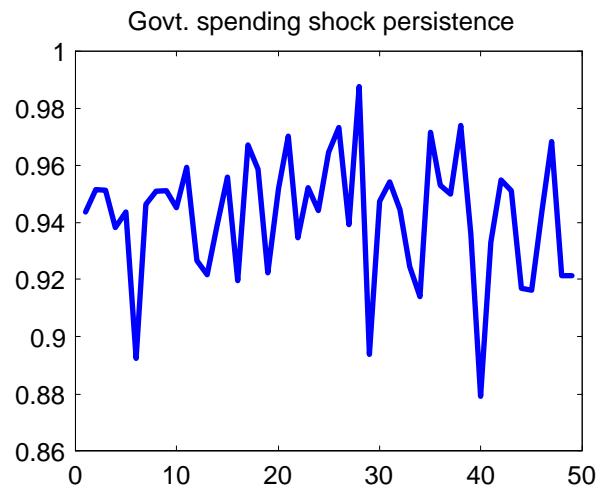
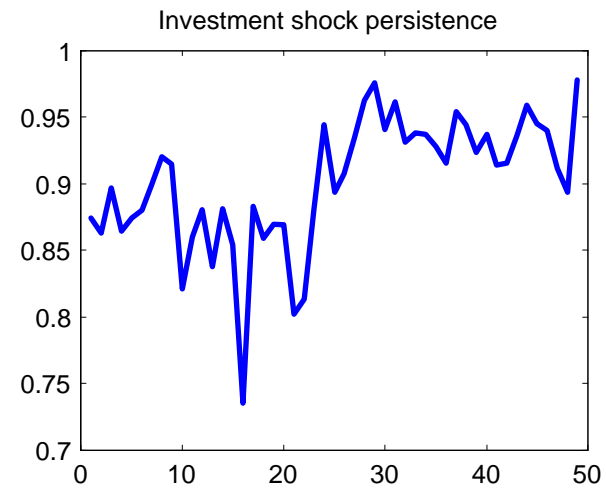
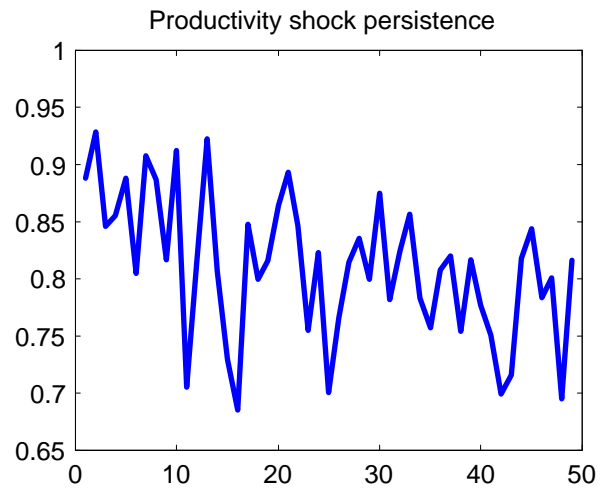
Application: SW's DSGE vs. BVAR

- SW sample: quarterly data 1970:2-1999:4 on DGP, consumption, investment, prices, real wages, employment, real interest rate
- Two questions:
 1. Is there evidence of parameter variation in the DSGE parameters?
 2. Was the DSGE *consistently* and *significantly* better than atheoretical Bayesian VARs over the sample?

Q#1: time-variation in DSGE parameters

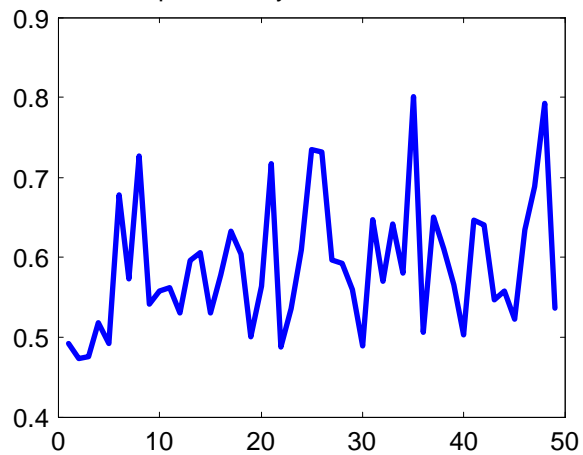
- Estimate DSGE model recursively using rolling samples of size $R = 70$ and plot posterior mode
 - persistence of shocks
 - standard deviation of shocks
 - monetary policy parameters

Parameter variation - shocks' persistence

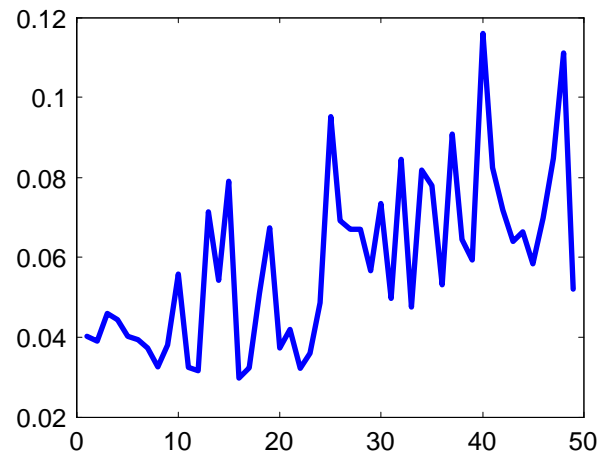


Parameter variation - standard deviation of shocks

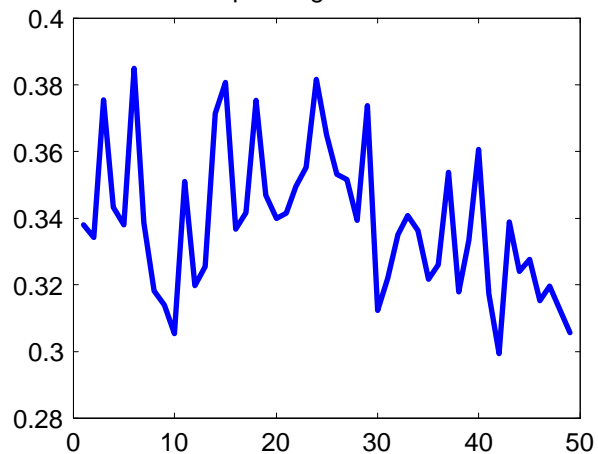
productivity shock st. dev.



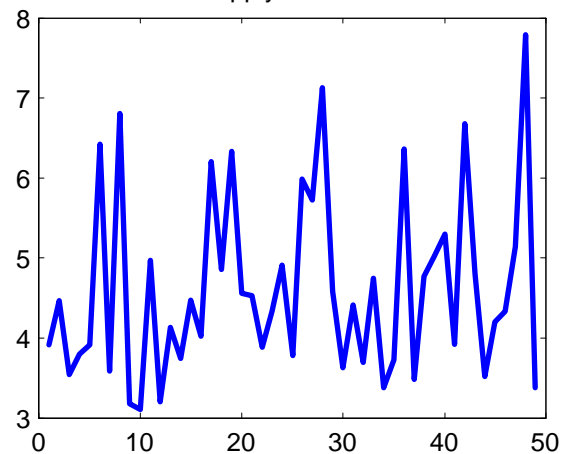
investment shock st. dev.



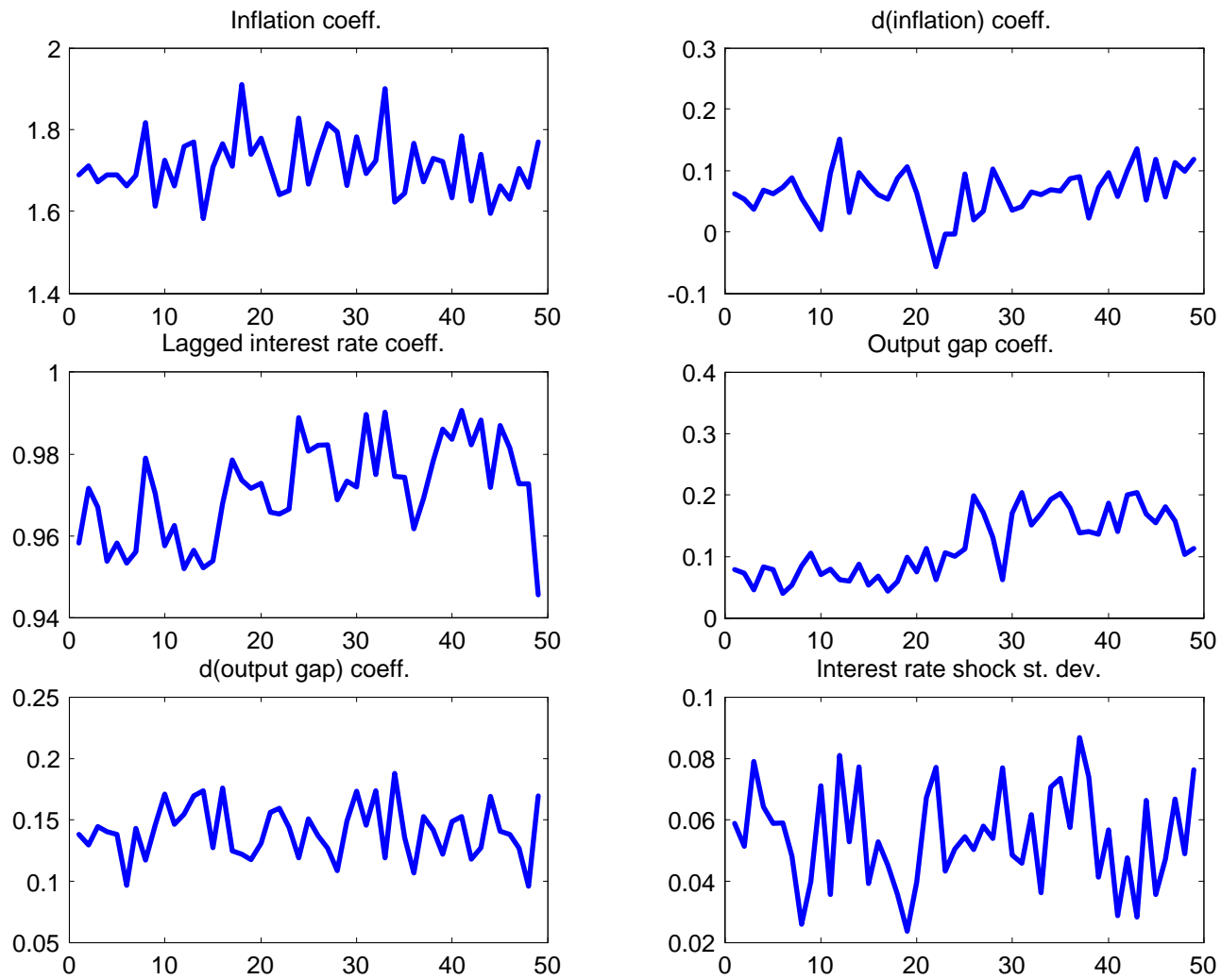
Govt. spending shock st. dev.



Labor supply shock st. dev.



Parameter variation – monetary policy parameters



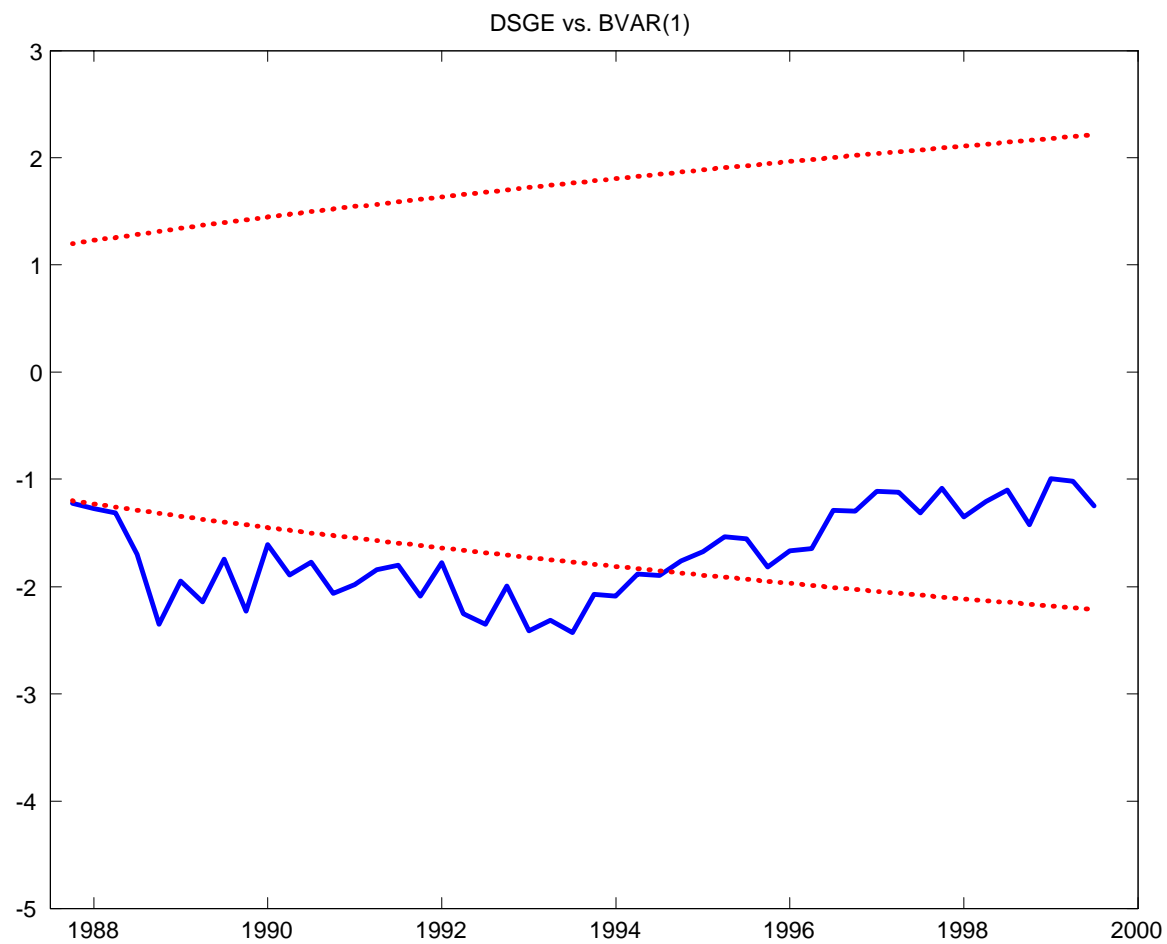
Summary: time-variation in structural parameters

- Moderate evidence of parameter variation in the DSGE model:
 - persistence of productivity shock ↓
 - persistence and standard deviation of investment shock ↑
 - monetary policy coefficients fairly stable

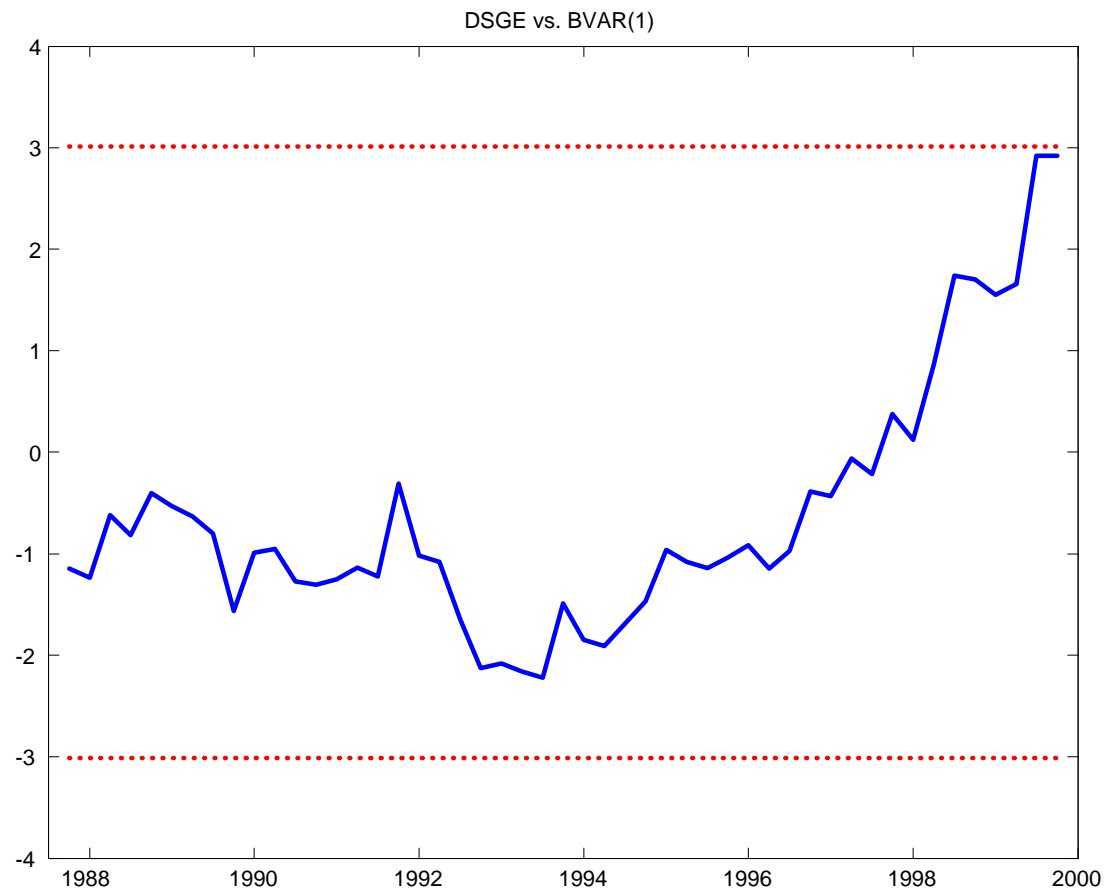
Q#2: time-varying performance of DSGE vs BVAR

- Recursive and rolling fluctuation tests. Total sample $T = 118$. Initial estimation sample $R = 70$
- DSGE vs. BVAR(1) with Minnesota priors
- Compute sequences of difference in average log-likelihoods $Q_t(\hat{\theta}_t) - Q_t(\hat{\gamma}_t)$ for $t = R, \dots, T$
- $\hat{\theta}_t$ is the posterior mode (a consistent estimator of θ_t^*)

Fluctuation test - recursive



Fluctuation test - rolling



Conclusion and extensions

- Proposed a formal method for evaluating time-variation in relative performance of misspecified non-nested models
- Empirical application confirmed SW's result that a DSGE has comparable performance to a BVAR in recent years
- Extension: optimal fluctuation test (against random walk alternative)