

Dynamic herding behavior in Pacific-Basin markets: Evidence and implications

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Abstract

This study examines investor herding behavior in Pacific-Basin equity markets. We find that herding is present in both rising and falling markets. Importantly, the level of herding is time-varying. We also find that herding is positively related to stock returns, but negatively related to market volatility. Herding estimates across markets are positively correlated, signifying co-movement of investor behavior in the region.

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1. Introduction

The term “herding” is often used in the finance literature to describe the correlation of investor behavior resulting from imitating other investors’ trading activity. This correlation in activity may stem from informational cascades, as the observation of prior trades can be so informative that investors are willing to ignore their own current private information in trading. As a result, herding behavior leads a group of investors to move in the same direction, pushing stock prices further away from their economic fundamentals, causing price momentum and excess volatility (Bikhchandani et al. 1992; Nofsinger and Sias 1999). Academic researchers in behavioral finance identify herding as the collective irrationality of investors, leading to the mispricing of economic fundamentals (Shiller, 2005).¹

To examine herding behavior, researchers commonly use a regression model to examine the cross-sectional dispersion of stock returns in response to changes in market conditions. Chang et al. (2000) find significant evidence of herding in South Korea and Taiwan and partial evidence of herding in Japan, but no evidence in the US and Hong Kong. Zhou and Lai (2009) focus on the Hong Kong equity market, and document that herding activity there tends to be more prevalent in small stocks. They also find that investors are more likely to herd when selling rather than buying stocks, meaning that herding activity appears to be more pronounced in falling markets vis-à-vis rising

¹This type of herding is referred to as *intentional* herding. In contrast, *unintentional* herding is mainly driven by fundamental (Bikhchandani and Sharma, 2001). It forms as investors receive correlated private information, share a similar educational background, and have equivalent analytical skills. As a result, they make similar investment decisions (see Hirshleifer, Subrahmanyam, and Titman, 1994).

markets. Demirer and Kutan (2006) investigate the Chinese market, and find no evidence of herding, suggesting that participants in the Chinese stock market make investment choices rationally. However, using a different data set, Tan et al. (2008) report that herding in Chinese stock markets occurs under both rising and falling market conditions. By analyzing industry-level data in a recent period, Chiang and Zheng (2010) document that herding exists in all Asian markets, but not in the US market. In summary, the evidence from international investors shows mixed results, and finds that herding behavior is most prevalent in emerging markets.

Although the aforementioned studies have provided insight regarding herding behavior in various aggregate markets, these empirical analyses have two limitations. First, their approach is essentially restricted to a single market analysis, and no attempt is made to detect the interaction of herding behavior across national borders. Our current study is motivated by recent empirical evidence that greater co-movement by investors often takes place during periods of high volatility. An increase in volatility tends to intensify herding activity. Boyer et al. (2006) and Chiang et al. (2007) document that contagion effects spread financial risk across markets, and herding activity further exacerbates market crises. The recent sub-prime crisis in the US market attests to the global nature of crisis transmission. Thus, models of herding behavior that do not consider cross-market repercussions may be subject to misspecification errors.²

The second limitation of previous studies of herding is that they use a constant coefficient model. The estimated coefficients essentially reflect an *average* value of a functional relation over a specific sample period, and herding behavior is assumed to be

² Chiang and Zheng (2010) investigate the cross-market correlation of stock return dispersions. However, correlation of return dispersions does not imply the correlation of herding coefficients. Thus, their analysis did not take into account the interdependence of herding behavior across different national stock markets.

unchanged throughout the period under investigation. In fact, when the market undergoes extreme stress, structural changes are likely to result. The constant coefficient estimators no longer provide updated information to reveal market dynamics. A time-varying coefficient model is needed to capture the changing nature of herding behavior.

To address the aforementioned empirical issues, this paper provides new evidence that helps to resolve the mixed findings of herding behavior in Pacific-Basin markets. More specifically, we use a Kalman-filter-based model that allows us to derive the time-varying nature of herding behavior. Our evidence shows that dynamic herding is significantly correlated with current stock returns and consistent with a positive-feedback trading hypothesis. However, the degree of herding becomes less pronounced as investors encounter uncertainty, whether this uncertainty arises from the domestic market, as measured by conditional variance, or from the global market, proxied by the implied volatility in the US market. The herding coefficients among different international investors are highly positively correlated after controlling for the common global volatility factor.

The remainder of this paper is organized as follows. Section 2 presents the procedures used to test herding behavior in previous studies. Section 3 describes the data. Section 4 reports empirical evidence of herding behavior using a constant coefficient model. Section 5 develops a time-varying coefficient model to estimate dynamic herding behavior. Section 6 examines the factors that explain herding dynamics. Section 7 presents additional evidence on herding by considering dynamic changes in state variables. Section 8 concludes.

2. Detecting herding behavior by investors

Conventional studies in the herding literature (Christie and Huang, 1995; Chang et al., 2000; Hwang and Salmon, 2004; Gleason, et al., 2004; Demirer and Kutan, 2006; and Tan et al., 2008) assume that investors act on common beliefs emerging from the market or follow observed momentum. Thus, market conditions and information on prevailing collective movements form the basis for guiding investment decisions. Christie and Huang (1995) and Chang et al. (2000) argue that during normal periods, rational asset pricing models predict that the dispersion in returns will increase with the absolute value of the market return, since individual investors are trading with their own private information, which is diverse. However, during periods of extreme market movements, investors tend to suppress their own beliefs and are likely to follow the market consensus. Christie and Huang (1995) propose that herding can be detected by examining the statistical significance of individual stock-return dispersions in response to indicator variables that capture the time periods of extreme returns.

Chang et al. (2000) propose that the formation of herds is more likely to be present during periods of relatively large price swings, and suggest that variations in individuals' sentiments regarding investment activity may be reflected in the dispersions of the cross-sectional stock returns. To measure this phenomenon, they propose the cross-sectional absolute deviation, $CSAD_t$, which is given by:

$$CSAD_t = \frac{1}{N} \sum_{i=1}^N |R_{i,t} - R_{m,t}| \quad (1)$$

where $CSAD_t$ is a measure of stock return dispersions, N is the number of firms (stocks) in the market or portfolio, and $R_{i,t}$ and $R_{m,t}$ are the value of a firm's realized return and

an equally weighted realized return of firms on day t . To conduct a test for detecting herding activity, previous studies have used the following model:

$$CSAD_t = \beta_0 + \beta_1 |R_{m,t}| + \beta_2 R_{m,t}^2 + \varepsilon_t \quad (2)$$

where $CSAD_t$ is a measure of cross-sectional stock return dispersions, and $R_{m,t}$ is the return on the market index on day t .³ Both the absolute value and squared value of $R_{m,t}$ appear on the right-hand side of Equation (2). This is based on the rationale that under normal conditions, a linear positive relationship between the return dispersion and market return is anticipated. However, during periods of relatively large price swings, in which market participants are more likely to follow the average consensus of all market opinions, the relation between $CSAD$ and the average market return is more likely to be nonlinear and negative. A significantly negative coefficient β_2 indicates the presence of herding behavior, since it reflects the phenomenon that during periods of market stress, return dispersions decline.

3. Data

The daily stock data employed in this study consist of individual firm stock prices and market price indices. The data cover five developed markets, including Australia (AU), Hong Kong (HK), Japan (JP), Singapore (SG), and the US, and six emerging Asian markets, including China (CN), Indonesia (ID), Malaysia (MA), South Korea (KR), Thailand (TH), and Taiwan (TW). The sample period is from 7/2/1997 through 3/23/2009. The exception is the US, for which the ending date is 12/31/2008. The choice

³ This equation has been used by Chang et al. (2000), Gleason, et al. (2004), Tan et al. (2008), and Chiang and Zheng (2010), among others. Alternatively, the cross-sectional squared deviation, $CSSD_t$, is also used in empirical tests of herding. However, as indicated by Duffee (2001), $CSSD_t$ is more sensitive to the outliers. For this reason, we use $CSAD_t$ in this study.

of starting date is constrained by the availability of data for the Chinese market, and the rationale for using 7/2/1997 is that this date is considered to be the beginning of the Asian financial crisis. Missing information for holidays is carefully inspected or interpolated. The stock return is calculated as the change in the natural logarithm of stock price (index) times 100. With the exception of the US data, which is taken from the *Compustat / CRSP* files, all of the data are taken from *Datastream International*. The plots of stock index series are shown in Figure 1.

<Figure 1 about here>

Table 1 provides summary statistics of *CSAD* in individual company stock returns for different markets spanning the period from 7/2/1997 to 3/23/2009.⁴ The cross-section consists of eleven markets, with the number of companies in each market ranging from 452 (Indonesia) to 6614 (the US). We find that the US has the highest mean value of return dispersion, followed by South Korea and Hong Kong.

Panel B of Table 1 presents the correlation matrix of *CSADs* among different markets. With the exception of the correlation for the *CSAD* between the China and Taiwan markets, which shows a negative sign, all of the other pairs of correlations are positive and highly significant. A special case is the Malaysian market, which has the highest correlation (0.593) with Singapore, and the lowest correlation (0.047) with Australia.⁵

[Table 1 about here]

⁴ The exception is the US data, for which the date range is from 7/2/1997 to 12/31/2008.

⁵ The high correlation between Malaysia and Singapore may be attributable to their traditional ties in political constitution and economic system before Singapore's separation from Malaysia in 1965. In addition, these two countries are geographically connected.

4. Evidence of herding behavior

4.1. Preliminary result

Equation (2) is estimated using a consistent estimator (Newey-West, 1987).⁶ As stated earlier, a negative value on the coefficient of $R_{m,t}^2$ is interpreted as the existence of herding. The evidence in Table 2 shows that the estimated equation for each market has high explanatory power, as the adjusted R -squared ranges from 0.29 to 0.67. Strikingly, all of the values of β_2 display a negative sign and are statistically significant at the 1% level. The results provide strong support for the existence of herding behavior in these markets. This finding is consistent with the multimarket analysis by Chiang and Zheng (2010). However, our evidence differs from the results reported by Chang et al. (2000), who find no evidence of herding in the US and HK markets, and only partial evidence of herding in the JP market.⁷ The result for the Chinese markets differs from the study reported by Demirer and Kutan (2006), who find no evidence of herding and conclude that participants in Chinese stock markets make rational investment choices.

[Table 2 about here]

4.2. Asymmetry of herding behavior

Recent empirical research (Bekaert and Wu, 2000) recognizes the significance of asymmetric responses of volatility to different outcomes of market shocks. Tan et al. (2008) and Chiang and Zheng (2010) provide evidence to substantiate asymmetric

⁶ Similar results are achieved by using a weighted least squares estimator.

⁷ Our data are somewhat different from those of Chiang and Zheng (2010), in that they use industry-level data to construct $CSAD_t$, while this paper uses individual firm-level data to construct $CSAD_t$. Our data are also different from the data used by Chang et al. (2000). They use daily firm returns over the period January 1963 to December 1997. Our data cover a more recent period, from 7/2/1997 to 3/23/2009. Chang et al. (2000) cover five markets: US, Hong Kong, Japan, South Korea, and Taiwan. Our data cover an additional six markets.

herding behavior under different market conditions. To examine whether herding behavior presents an asymmetric reaction on days when the market is rising vis-à-vis days when the market is falling, we rewrite Equation (2) conditional on the following two regimes:

$$\begin{aligned}
 CSAD_t^+ &= \beta_{01} + \beta_{11}|R_{m,t}^+| + \beta_{21}(R_{m,t}^+)^2 + \varepsilon_{1,t}, \text{ if } R_{m,t} > 0 \\
 CSAD_t^- &= \beta_{02} + \beta_{12}|R_{m,t}^-| + \beta_{22}(R_{m,t}^-)^2 + \varepsilon_{2,t}, \text{ if } R_{m,t} < 0
 \end{aligned} \tag{3}$$

where $R_{m,t}^+$ ($R_{m,t}^-$) is the equally weighted market return during period t when the market is rising, denoted by “+” (falling, which is denoted by “-“), and the second subscripts 1 and 2 on the β coefficient identify rising and falling regimes. Table 3 presents the estimated herding coefficients under rising (Panel A) and falling (Panel B) market conditions. We continue to find a statistically significant negative sign for the herding coefficient, regardless of whether the market is rising or falling, implying that herding occurs in either regime.

[Table 3 about here]

To test the asymmetry of herding coefficients between rising and falling markets, we conduct a Wald test by subtracting the coefficient on the falling markets from the coefficient on the rising markets, $(\beta_{21} - \beta_{22})$. The results are reported in Panel C of Table 3, and indicate that asymmetry is significant in 9 out of 11 markets. The exceptions are Singapore and Thailand. In the majority of markets, the herding effect appears to be stronger during rising markets than during falling markets.

One possible explanation for an asymmetry in herding between rising and falling markets is the flow of positive and negative information. For example, if analysts tend to issue buy recommendations more actively than sell recommendations, and investors act on the recommendations, we would observe more herding in rising markets. Another possibility is that investors believe that the government will intervene when markets decline significantly, and are thus less likely to herd in falling markets.

5. Analysis of time-varying herding behavior

5.1 The model

The conventional estimation using a constant regression model in the above analysis provides evidence on the *average* relation over time between stock return dispersions and squared market returns. In other words, the estimated result for detecting herding is essentially static in nature. To capture the dynamic behavior of herding activity, we use a Kalman-filter-based model, which is expressed as:

$$CSAD_t = \beta_{0,t} + \beta_{1,t}|R_{m,t}| + \beta_{2,t}R_{m,t}^2 + \varepsilon_t \quad (4)$$

$$\beta_{i,t} = \beta_{i,t-1} + v_{i,t}, \quad v_{i,t} \sim N(0, \sigma_{v,i}^2), \quad \text{where } i=0,1, \text{ and } 2. \quad (5)$$

Equation (4) is the measurement equation and $[\beta_{0,t}, \beta_{1,t}, \beta_{2,t}]'$ is a vector of state variables.

Equation (5) is called a transition equation, and the state variables are assumed to evolve with a random walk process. A special feature of this model is that the coefficients are time-varying and can be obtained by using the Kalman filter procedure.⁸

⁸ One can substitute equation (5) into equation (4) to obtain an expanded regression. Maddala and Kim (2002) first use an OLS method to obtain the estimated variances of ε and v . Then, they apply GLS in the estimation. Harvey (1990), Hamilton (1994), and Tsay (2005) provide a detailed description of the Kalman filter estimation.

5.2. Estimates of time-varying herding behavior

Figure 2 presents the plots of the time series estimates of herding coefficients for 11 markets based on the state-space model. The herding series are time-varying and appear to be stationary. Most of the markets exhibit coefficients around the -0.2 level. The exception is the herding coefficient in the US market, which most of the time lies in the positive region, showing an absence of herding activity. This finding is in contrast to the result based on the constant regression estimation of herding in the US market, indicating a misleading statistical inference.

<Figure 2 about here>

To gain more insight into the time-varying nature of herding, we present summary statistics and the correlation matrix of the herding coefficients in Tables 4 and 5, respectively. Several interesting points emerge. The mean and median values for all of the non-US markets are negative, coupled with the statistics of skewness to be distributed to the left tail, suggesting that herding activity is widely present in these regions. This finding is consistent with the evidence reported in Table 2, except for the US market.

Note that the constant coefficient estimated using OLS is not an average of the time-varying coefficients, as noted by Garbade and Wachtel (1978). For instance, the average slope coefficients for AU, HK, and JP are -0.259, -0.296, and -0.267 from the time-varying coefficient model versus -0.060, -0.058, and -0.057 for the constant coefficient model, respectively. More dramatically, in the US market, the mean value is 0.04 in the

time-varying coefficient model versus -0.021 in the constant coefficient estimator. The time-varying coefficient evidence is consistent with the earlier finding for rejecting the null hypothesis of equal coefficients during rising and falling markets, as indicated by the Wald test. The evidence in Table 4, in conjunction with the time series plots of Figure 2, clearly indicates that herding activity varies over time, suggesting that using a constant coefficient approach cannot accurately test for herding behavior.

[Table 4 about here]

Statistics of maximum value for some markets, such as AU, HK, JP, and TH, exhibit a negative sign, meaning that herding behavior is consistently and unambiguously present throughout the sample period under investigation. The max-min range and standard deviation of the herding series indicate that HK, ID, AU, and TH exhibit much higher variations than other markets.

Table 5 reports the correlation matrix of herding coefficients. It is notable that, in general, herding in all the markets in the Pacific-Basin is highly correlated in a positive direction.⁹ Except for Taiwan, herding in all countries is negatively correlated with herding in the US market. The evidence is consistent with the earlier finding by Chiang et al. (2007) in their study of the Asian crisis. However, our results suggest that the co-movement of herding extends beyond the crisis period. Consistently, the negative correlations of the herding series for most Pacific-Basin markets with the US market indicate a lack of co-movement of herding with the US market, although the stock return dispersions are positively correlated, as shown in Table 1. Correlation in stock return dispersions does not imply correlation in herding. The former reflects the relation of

⁹ There are a few cases that show a negative correlation with the series for Indonesian and Malaysian markets.

cross-sectional return dispersions between countries, market risk (Chiang and Zheng, 2010), while correlation of herding behavior (denoted by the $\beta_{2,t}$ coefficient) reflects how the return dispersions commonly react to market conditions in the two countries.

[Table 5 about here]

6. Determinants of herding dynamics

6.1 Domestic market factors

Now that we have empirically detected that herding behavior is time-varying for a group of markets in the Pacific Basin. The next job is to identify its determinants. To address this issue, we form the hypotheses below.

(i) *Stock performance hypothesis*: Herding activity is perceived to be correlated with stock market performance. Research demonstrates that investors react to stock prices (Grinblatt et al., 1995). Evidence (DeLong et al., 1990; Sentana and Wadhvani, 1992) suggests that positive feedback traders buy stocks in a rising market and sell stocks in a falling market, while negative feedback traders follow an investment strategy of “buy low and sell high.” Theoretically, there is no *a priori* reason to believe which type of feedback-trading strategy herding investors intend to follow, so we are unable to impose a definite sign on the coefficient on the stock market return.

(ii) *Volatility hypothesis*: Herding behavior is likely to be related to stock market volatility. This hypothesis is associated with the fundamental characteristics of herding formation that occur during periods of market stress. The empirical literature (Butler and Joaquin, 2002; Forbes and Rigobon, 2002; Corsetti et al., 2005) shows that cross-market correlations rise in periods of high volatility. These studies provide us with the

motivation to test whether herding activity is influenced by stock return volatility. Following the conventional approach, we model the conditional stock return variance as an asymmetric GARCH(1,1)-type process.

To examine the above hypotheses, we write the following regression model:

$$Herd_t = \phi_0 + \phi_1 R_{m,t} + \phi_2 \hat{\sigma}_t^2 + \varepsilon_t \quad (6)$$

where the dependent variable, $Herd_t$, is the estimated herding coefficient on day t , $\beta_{2,t}$, which is derived from the state-space model in equation (4); $R_{m,t}$ is stock market return, $\hat{\sigma}_t^2$ is the conditional variance of stock returns based on asymmetric GARCH(1,1) process for each market, ϕ_i ($i= 0,1, 2$) denotes constant parameters, and ε_t is an error term.¹⁰

Table 6 reports consistent estimates for each hypothesis involving univariate and multivariate regressions. To avoid possible disturbance due to the Asian crisis, we use the sample period spanning from January 4, 1999 to March 23, 2009. Several empirical findings are worth noting. First, the estimated coefficients for stock market returns show a negative sign, and the majority of the coefficients are statistically significant.¹¹ Since the herding coefficient, β_2 , is negative, the inverse relation between the herding coefficient and stock returns is consistent with a market phenomenon that when stock

¹⁰ Alternatively, we also measured stock returns by using recent momentum based on the past five days' moving average return for an individual stock market, i.e., $(1/5) \sum_{s=1}^5 R_{m,t-s}$. The results are similar to those reported in Table 6.

¹¹ Less clear are the cases for Malaysia and Singapore. Note that the Malaysian financial system is mainly governed by *Shariah* law. The IMF Survey Magazine (2007) reports that *Shariah*-compliant stocks comprise about 86 percent of domestic stock exchange. The Survey indicates that the Ninth Malaysia Plan allocates 200 billion Malaysian ringgit for developmental projects and 20 billion for private finance initiatives. The system in Singapore is highly tied to Malaysia as we observe in footnote 5.

returns rise, the observed herding measure increases. Our interpretation is that when stock markets experience positive returns, investors tend to discard their own private beliefs and move along with market consensus, and thus herding increases.¹² However, as markets suffer losses, investors may be less likely to act in a coordinated manner, since they are unwilling to assume immediate losses and hesitate to sell their shares as stock prices fall (Statman et al., 2006). The evidence is in agreement with our earlier result that herding activity is more profound in a rising market than in a falling market. Our finding is also consistent with the finding provided by Duffee (2001) in that “aggregate trading volume tends to be higher on days when the stock market rises than when it falls.”

[Table 6 about here]

Herding is positively correlated with stock return volatility, as all of the estimated slopes on the conditional variance variable are statistically significant at the 1% level. Since the herding measure itself has a negative value, a positive sign of the estimated coefficient means that the magnitude of the measured herding coefficient declines as conditional volatility increases. This is consistent with a market phenomenon that as uncertainty increases, investors who lack clear market signals and fundamentals are prone to act independently.

The evidence presented in Table 6 suggests that investor reaction to the market state variables displays a complicated nonlinear behavior. Herding behavior is not only influenced by the occurrence of stressful events, but is also sensitive to prevailing market returns and conditional variance, among other factors. It appears that herding behavior is

¹² Our evidence here is not contradictory to the finding reported in Table 3, where we examine how stock return dispersions react to extreme changes in market returns during rising and falling markets. Here we interpret herding behavior in response to recent market conditions.

rather complex, and the evidence shows that it is consistent with a positive feedback trading process.

6.2. *Cross-market factors*

Evidence presented in the previous section focuses only on the domestic factors affecting herding dynamics. In this section, we examine whether herding in a given market is affected by herding in other markets and global stock market volatility. The correlation in herding may be due to geographic proximity that produces close trading relation in the region, or to a similar cultural background with less transparency and less public information available, which would induce investors to form a correlated trading decision. The experience from the Asian crisis period indicates that herding behavior tends to display co-movements (Marais and Bates, 2006; Chiang et al., 2007). Our earlier evidence in Table 5 also indicates that the markets in the Pacific Basin exhibit a high degree of correlation. This evidence motivates the inclusion of a cross-market herding term in our model. The criterion for the selected cross-market factor is based on the simple correlation from Table 5, the visibility of the capital market, the region's dominance in trades, or the distance of the geographic location with respect to the local market.

The second cross-market factor is global market volatility, referred to as interdependence or spillovers from stock return volatility (Diebold and Yilmaz, 2009). Experience in recent financial crises indicates that it does not matter through which channel the volatility is transmitted (Pericoli and Sbracia, 2003).¹³ Whenever negative news develops in a given market, it will soon be learned by participants in other markets.

¹³ Pericoli and Sbracia (2003) provide a theoretical framework to highlight different channels for the international transmission of financial shocks. They show how crises that occur in one country can be transmitted across countries, without assuming ad hoc portfolio management rules or market imperfections.

As a result, international investors will reallocate their portfolios to hedge the shocks. Ng (2000) and Beirne et al. (2009) find evidence of significant stock-return volatility spillovers from the US market to many Pacific-Basin countries.

To assess the impact of the volatility spillovers, we employ the Chicago Board Options Exchange (CBOE) Volatility Index, commonly referred to as the *VIX*. The *VIX* is constructed using implied volatilities from options, and represents a forward-looking measure of the expected return volatility on the S&P 500 index over the next 30 days. Whaley (2009) observes that the *VIX* spikes during periods of market turmoil, reflecting market conditions that if expected market volatility increases (decreases), stock prices fall (rise), and investors require a higher (lower) rate of return to compensate for bearing risk. Connolly et al. (2005) find that the *VIX* can be considered a primary measure of perceived stock market risk or uncertainty. Hakkio and Keeton (2009) further argue that the *VIX* can capture uncertainty arising from asset fundamentals or unexpected shifts in investor sentiment. Masih and Masih (2001) note that the financial information in the US market is considered to be most influential, and the *VIX* is also widely used in the literature (Beirne et al., 2009). Empirically, it is of interest to examine whether the implied volatility of the US stock market (*VIX*) would have an effect on local stock volatility and herding behavior. For this reason, we modify equation (6) as:

$$Herd_t = \phi_0 + \phi_1 R_{m,t} + \phi_2 \hat{\sigma}_t^2 + \phi_3 (Cross\ HERD)_t + \phi_4 VIX_{t-1} + \varepsilon_t \quad (7)$$

where the variable *Cross HERD*_{*t*} is the herding coefficient from the country that exhibits the greatest correlation with the herding coefficient that is the dependent variable.

Table 7 reports the coefficient estimates with cross-market herding (*Cross HERD*)_{*t*} and the lagged *VIX* as incremental variables. Including these additional

variables significantly improves the explanatory power of the equation. The adjusted R -squared values in Table 7 range from 0.16 to 0.54, compared with values of 0.08 to 0.24 in the models in Table 6 that did not include these additional control variables. However, the qualitative results of the market return and conditional stock-return variance remain significant, as we found earlier. Interestingly, the cross-market herding coefficient is positive and statistically significant at the 1% level for all markets. This finding suggests that investors' herding behavior in this region exhibits co-movement.¹⁴

Turning to the coefficients on VIX_{t-1} , we find evidence that with the exception of two advanced markets, Australia and Japan, all of the estimated coefficients in other Asian markets are positive and significant, suggesting that greater stock market volatility in S&P 500 returns can influence herding behavior in the local market. The impact of the lagged VIX variable on herding is likely to carry an effect similar to that of local conditional volatility.¹⁵

Based on the evidence from the literature (Beirne et al., 2009), we anticipate that $\hat{\sigma}_t^2$ and VIX_{t-1} may be positively correlated with each other due to either the cross-listing of shares (Howe and Madura, 1990) or a contagion effect (King and Wadhvani, 1990). Table 8 reports the correlation analysis between the lagged VIX_{t-1} and conditional variance for each market. As can be seen in the first column, the correlation coefficients range from 0.16 to 0.50, with the highest value in South Korea. This is consistent with the finding in Table 7 vis-à-vis the evidence in Table 6, where the coefficient on $\hat{\sigma}_t^2$ is

¹⁴ This evidence and the correlation analysis in Table 5 are consistent with the results of Chiang and Zheng (2010), who find that stock return dispersions are correlated across international markets, but do not test for the interdependence of herding behavior across different markets.

¹⁵ Using the current value of the VIX in the estimations, we find the results are comparable. Since the trading dates in Asian markets are about 12 hours ahead of the US market, using a lagged VIX is more relevant.

insignificant when the VIX_{t-1} is included in the estimated equation. Although the evidence shows that $\hat{\sigma}_t^2$ and VIX_{t-1} to some extent are correlated, both $\hat{\sigma}_t^2$ and VIX_{t-1} are still individually significant, meaning that each variable plays a role in explaining investors' behavior. In particular, $\hat{\sigma}_t^2$ represents local stock market risk, while VIX_{t-1} corresponds to global market risk. Both variables have a similar qualitative impact on herding.

[Table 7 about here]

7. Implications of dynamic herding behavior

The original herding model proposed by Chang et al. (2000) in Equation (2) posits that herding exists if stock return dispersions negatively respond to extreme market conditions as measured by the squared market return. The evidence from the state-space model clearly indicates that herding behavior is time-varying and sensitive to the prevailing market returns and conditional return volatility. This finding implies that investors' behavior when facing stressful market conditions is rather complex and the nonlinear reaction cannot be adequately captured by a simple term represented by $R_{m,t}^2$. To illustrate this point and simplify the analysis, we incorporate just the domestic arguments of Equation (6) into the original equation proposed by Chang et al. (2000). Doing so yields the following reduced-form equation:

$$CSAD_t = b_0 + b_1 |R_{m,t}| + b_2 R_{m,t}^2 + b_3 R_{m,t}^3 + b_4 \hat{\sigma}_t^2 R_{m,t}^2 + \varepsilon_t \quad (8)$$

Comparing equation (8) with (2) indicates that two extra terms, $R_{m,t}^3$ and $\hat{\sigma}_t^2 R_{m,t}^2$, are excluded from Chang et al.'s (2000) original equation. In particular, $R_{m,t}^3$ can be viewed as a product term of $R_{m,t}^2$ and $R_{m,t}$, which represents the interaction of herding behavior

with stock returns, and the product term $\hat{\sigma}_t^2 R_{m,t}^2$ captures the interaction of herding with conditional variance. Equation (8) thus contains multiple elements of nonlinearity, reflecting various reactions of stock return dispersions to extreme market conditions.

Table 9 reports the estimates of the augmented equation. The incremental variables, $R_{m,t}^3$ and $R_{m,t}^2 \cdot \hat{\sigma}_t^2$, are statistically significant, especially the $R_{m,t}^2 \cdot \hat{\sigma}_t^2$ term. The F-statistics show that the joint test is rejected for all of the markets, suggesting that the state variables provide additional information to explain cross-sectional dispersion in returns.

Since the inclusion of the incremental variables is likely to change the magnitude of the standard measure of herding, we test whether the herding coefficient, b_2 , of the augmented equation is significantly different from the coefficient, β_2 , in Chang et al.'s (2000) original model. The Chi-squared statistics in Table 9 clearly indicate that except for the South Korean market, the null hypothesis, $(b_2 - \beta_2) = 0$, is rejected with a negative sign. These results suggest that using the augmented equation in estimation contains richer information about market conditions and herding behavior. On average, the coefficient on the squared market return is more negative in the augmented specification than in the original one, suggesting that the original specification may be less able to detect the presence of herding.

8. Summary and Conclusions

This study examines investor herding behavior for ten Pacific-Basin markets: Australia, China, Hong Kong, Indonesia, Japan, Malaysia, South Korea, Singapore, Thailand, and Taiwan. To incorporate a global influence, we also include US data. By applying a constant coefficient regression model using daily data from July 2, 1997

through March 23, 2009 for individual firm stock returns, we find significant evidence of herding in each national market, including the US. This result stands in contrast to the earlier evidence of no herding in advanced markets (Chang et al. 2000) or in Chinese markets (Demirer and Kutan, 2006).

The conventional estimation approach using the constant coefficient regression model provides information on the *average* relation between stock return dispersion and extreme movements in stock prices. Hence, the estimated results for detecting herding are essentially static in nature. We adopt a time-varying approach by using a Kalman-filter-based model to estimate dynamic herding behavior. The evidence shows that herding behavior for the markets under investigation is time-varying. The exception is the US market, which shows a positive sign.

This study shows that dynamic herding behavior is significantly correlated with state variables such as current stock returns, conditional stock-return variance, and global stock volatility (*VIX*). The evidence suggests that an increase in stock returns leads to an increase in the herding measure, while a rise in local stock-return volatility or the *VIX* tends to reduce observed herding activity. We find that herding coefficients are positively correlated across markets, suggesting the interdependence of herding behavior in the Pacific-Basin region. An important implication emerging from this study is that investors' reaction to extreme market conditions is highly nonlinear. The evidence shows that herding behavior reacts not only to the occurrence of large swings in market prices, but also to the state of market return and volatility conditions. Omission of these state influences in an empirical model is likely to result in inaccurate inferences regarding herding.

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Table 1. Descriptive statistics of cross-sectional absolute deviations

This table lists descriptive statistics of daily, equally weighted cross-sectional absolute deviations ($CSAD_t$) for 5 developed markets, including Australia (AU), Hong Kong (HK), Japan (JP), Singapore (SG), and the US; 6 emerging Asian markets, including China (CN), Indonesia (ID), Malaysia (MA), S. Korea (KR), Thailand (TH), and Taiwan (TW). The data in this study are daily observations from individual firms for each market, spanning from 7/2/1997 to 3/23/2009. The exception is the US; the ending date of the US data is 12/31/2008. The starting period is constrained by the data available from the Chinese market and the date of 7/2/1997 is considered to be a breakout point for the Asian crisis. Missing information for holidays is carefully inspected or interpolated. Calculations of $CSAD$ are given by equation (1) as:

$$CSAD_t = \frac{1}{N} \sum_{i=1}^N |R_{i,t} - R_{m,t}|.$$

Panel A. Statistics of $CSAD$

| | Mean | Median | Max. | Min. | Std. dev. | Skewness | Kurtosis | No. of Firms | Obs |
|-------------|-------|--------|-------|-------|-----------|----------|----------|--------------|------|
| $CSAD_{AU}$ | 1.470 | 1.441 | 3.935 | 0.483 | 0.401 | -0.362 | 9.326 | 2360 | 3059 |
| $CSAD_{CN}$ | 1.385 | 1.372 | 4.180 | 0.497 | 0.632 | 0.060 | 4.035 | 863 | 3059 |
| $CSAD_{HK}$ | 1.775 | 1.772 | 4.657 | 0.809 | 0.644 | -0.403 | 5.135 | 1164 | 3059 |
| $CSAD_{ID}$ | 1.596 | 1.474 | 4.945 | 0.089 | 0.810 | 0.624 | 4.177 | 452 | 3059 |
| $CSAD_{JP}$ | 1.435 | 1.476 | 2.921 | 0.722 | 0.496 | -0.962 | 4.953 | 2918 | 3059 |
| $CSAD_{KR}$ | 2.016 | 1.765 | 5.851 | 0.300 | 0.910 | 0.533 | 3.881 | 976 | 3059 |
| $CSAD_{MA}$ | 1.586 | 1.514 | 3.967 | 0.850 | 0.648 | 0.129 | 4.437 | 986 | 3059 |
| $CSAD_{SG}$ | 1.631 | 1.567 | 4.185 | 0.821 | 0.578 | 0.143 | 5.263 | 707 | 3059 |
| $CSAD_{TH}$ | 1.149 | 1.080 | 4.447 | 0.559 | 0.517 | 0.609 | 5.818 | 693 | 3059 |
| $CSAD_{TW}$ | 1.609 | 1.609 | 3.648 | 0.103 | 0.586 | -0.629 | 4.558 | 849 | 3059 |
| $CSAD_{US}$ | 2.101 | 2.068 | 4.954 | 0.979 | 0.589 | 0.602 | 3.162 | 6614 | 2894 |

Table 1 (continued)

Panel B. Correlation matrix of the CSAD: 7/2/1997– 12/31/2008

This table contains the correlation matrix for the CSAD. Each column contains three statistics for each individual market. The first figure is the covariance, the second value is the correlation coefficient, and the last one is the t-statistic for the correlation coefficient. The data in this study span from 7/2/1997 to 3/23/2009. The exception is the US; the ending date of the US data is 12/31/2008.

| Markets | $CSAD_{AU}$ | $CSAD_{CN}$ | $CSAD_{HK}$ | $CSAD_{ID}$ | $CSAD_{JP}$ | $CSAD_{KR}$ | $CSAD_{MA}$ | $CSAD_{SG}$ | $CSAD_{TH}$ | $CSAD_{TW}$ | $CSAD_{US}$ |
|-------------|------------------------------|--------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|------------------------------|---------------------------|
| $CSAD_{AU}$ | 0.1422 1.0000 ----- | | | | | | | | | | |
| $CSAD_{CN}$ | 0.0533 0.2259 (12.469) | 0.3914 1.0000 ----- | | | | | | | | | |
| $CSAD_{HK}$ | 0.0860 0.3667 (21.192) | 0.1060 0.2723 (15.219) | 0.3867 1.0000 ----- | | | | | | | | |
| $CSAD_{ID}$ | 0.0335 0.1114 (6.029) | 0.0533 0.1070 (5.786) | 0.1631 0.3292 (18.747) | 0.6343 1.0000 ----- | | | | | | | |
| $CSAD_{JP}$ | 0.0286 0.1546 (8.416) | 0.0172 0.0560 (3.018) | 0.0766 0.2510 (13.944) | 0.1029 0.2633 (14.678) | 0.2406 1.0000 ----- | | | | | | |
| $CSAD_{KR}$ | 0.0237 0.0695 (3.747) | 0.0887 0.1567 (8.533) | 0.2295 0.4079 (24.020) | 0.3131 0.4345 (25.942) | 0.1743 0.3928 (22.965) | 0.8187 1.0000 ----- | | | | | |
| $CSAD_{MA}$ | 0.0112 0.0471 (2.533) | 0.0734 0.1863 (10.194) | 0.1318 0.3367 (19.227) | 0.2210 0.4407 (26.394) | 0.0583 0.1888 (10.339) | 0.2586 0.4539 (27.388) | 0.3964 1.0000 ----- | | | | |
| $CSAD_{SG}$ | 0.0311 0.1494 (8.124) | 0.0583 0.1687 (9.204) | 0.1598 0.4654 (28.272) | 0.1912 0.4347 (25.951) | 0.0643 0.2373 (13.137) | 0.2372 0.4746 (28.994) | 0.2060 0.5925 (39.544) | 0.3050 1.0000 ----- | | | |
| $CSAD_{TH}$ | 0.0205 0.1064 (5.751) | 0.0357 0.1117 (6.044) | 0.0970 0.3048 (17.207) | 0.1409 0.3459 (19.823) | 0.0671 0.2676 (14.932) | 0.1768 0.3821 (22.233) | 0.0956 0.2969 (16.719) | 0.0905 0.3205 (18.195) | 0.2616 1.0000 ----- | | |
| $CSAD_{TW}$ | 0.0477 0.2146 (11.816) | -0.0224 -0.0608 (-3.276) | 0.1352 0.3693 (21.364) | 0.1388 0.2960 (16.664) | 0.1377 0.4767 (29.158) | 0.2515 0.4720 (28.790) | 0.0816 0.2201 (12.133) | 0.1226 0.3769 (21.882) | 0.0718 0.2385 (13.207) | 0.3467 1.0000 ----- | |
| $CSAD_{US}$ | 0.0266 0.1226 (6.640) | 0.0305 0.0845 (4.561) | 0.1145 0.3193 (18.114) | 0.0735 0.1600 (8.714) | 0.0544 0.1924 (10.540) | 0.1458 0.2795 (15.652) | 0.0560 0.1544 (8.401) | 0.0706 0.2218 (12.233) | 0.0343 0.1163 (6.293) | 0.1262 0.3718 (21.536) | 0.3324 1.0000 ----- |

Table 2. Estimates of herding equation based on a constant coefficient model

This table reports the regression results of CSAD. The estimated equation is specified as:

$$CSAD_t = \beta_0 + \beta_1 |R_{m,t}| + \beta_2 R_{m,t}^2 + \varepsilon_t \quad (2)$$

where $CSAD_t$ is equally weighted cross-sectional absolute deviations; $R_{m,t}$ is the value of an equally weighted realized return of all firms' indexes on day t ; and $R_{m,t}^2$ is the squared term of $R_{m,t}$. AU, CN, HK, ID, JP, KR, MA, SG, TH, TW, and the US denote the markets for Australia, China, Hong Kong, Indonesia, Japan, South Korea, Malaysia, Singapore, Thailand, Taiwan, and the United States, respectively. The data in this study are daily observations from individual firms for each market, spanning from 7/2/1997 to 3/23/2009. The exception is the US; the ending date of the US data is 12/31/2008. \bar{R}^2 is the adjusted R -squared. The numbers in parentheses are t-statistics. ***, **, * denotes that the coefficient is significant at the 1%, 5%, and 10% levels, respectively.

| <i>Markets</i> | <i>C</i> | $ R_{m,t} $ | $R_{m,t}^2$ | \bar{R}^2 |
|----------------|---------------------|---------------------|-----------------------|-------------|
| Australia | 1.220 (88.86)*** | 0.782 (31.78)*** | -0.060 (-18.77)*** | 0.49 |
| China | 0.982 (29.32)*** | 0.417 (15.24)*** | -0.030 (-7.34)*** | 0.31 |
| Hong Kong | 1.283 (61.72)*** | 0.759 (29.39)*** | -0.058 (-10.66)*** | 0.56 |
| Indonesia | 0.939 (43.34)*** | 1.032 (36.36)*** | -0.084 (-16.59)*** | 0.66 |
| Japan | 1.086 (46.48)*** | 0.588 (16.97)*** | -0.057 (-5.37)*** | 0.36 |
| South Korea | 1.224 (40.54)*** | 0.816 (23.79)*** | -0.061 (-10.10)*** | 0.54 |
| Malaysia | 1.099 (61.55)*** | 0.587 (28.23)*** | -0.031 (-8.76)*** | 0.67 |
| Singapore | 1.122 (69.66)*** | 0.737 (39.79)*** | -0.059 (-16.60)*** | 0.66 |
| Thailand | 0.716 (46.71)*** | 0.896 (26.53)*** | -0.097 (-9.33)*** | 0.60 |
| Taiwan | 1.147 (40.75)*** | 0.643 (18.06)*** | -0.084 (-9.47)*** | 0.38 |
| United States | 1.769 (54.57)*** | 0.470 (14.33)*** | -0.021 (-4.56)*** | 0.29 |

Table 3. Estimates of herding equation in rising and falling stock markets

This table reports the regressions results of the herding equation for the United States and the Pacific-Basin stock markets in rising and falling markets. The estimated equations are:

$$CSAD_t^+ = \beta_{01} + \beta_{11}|R_{m,t}^+| + \beta_{21}(R_{m,t}^+)^2 + \varepsilon_{1,t}, \text{ if } R_{m,t} > 0$$

$$CSAD_t^- = \beta_{02} + \beta_{12}|R_{m,t}^-| + \beta_{22}(R_{m,t}^-)^2 + \varepsilon_{2,t}, \text{ if } R_{m,t} < 0$$

where $R_{m,t}^+$ ($R_{m,t}^-$) is the equally weighted market return during period t when the market is up, denoted by “+” (down, which is denoted by “-“). Panel A reports the estimates in rising markets, Panel B, for down markets, and Panel C reports the test statistics for the difference in herding coefficients between the two regimes. The sample period is from 7/2/1997 to 3/23/2009. The US data are from 7/2/1997 to 12/31/2008 Numbers in parentheses are t-statistics based on Newey-West (1987) consistent standard errors. ***, **, and * denote statistical significance at the 1%, 5%, and 10% levels, respectively.

Panel A: Regression results when market rises ($R_{m,t} > 0$)

| <i>Market</i> | <i>C</i> | $ R_{m,t} $ | $R_{m,t}^2$ | \bar{R}^2 |
|---------------|---------------------|---------------------|-----------------------|-------------|
| Australia | 0.382 (34.64)*** | 2.703 (58.81)*** | -0.388 (-29.37)*** | 0.55 |
| China | 0.291 (25.05)*** | 0.964 (53.53)*** | -0.116 (-30.72)*** | 0.55 |
| Hong Kong | 0.358 (29.49)*** | 2.089 (70.42)*** | -0.311 (-36.75)*** | 0.69 |
| Indonesia | 0.291 (27.59)*** | 1.768 (79.24)*** | -0.186 (-41.23)*** | 0.75 |
| Japan | 0.313 (29.20)*** | 1.471 (63.76)*** | -0.165 (-27.40)*** | 0.61 |
| South Korea | 0.384 (27.77)*** | 1.575 (66.90)*** | -0.159 (-34.66)*** | 0.68 |
| Malaysia | 0.344 (31.86)*** | 1.054 (72.96)*** | -0.061 (-40.93)*** | 0.68 |
| Singapore | 0.302 (29.09)*** | 1.612 (74.09)*** | -0.189 (-38.25)*** | 0.72 |
| Thailand | 0.215 (28.69)*** | 1.532 (75.57)*** | -0.207 (-35.49)*** | 0.71 |
| Taiwan | 0.307 (26.28)*** | 1.547 (65.90)*** | -0.248 (-38.53)*** | 0.65 |
| United States | 0.496 (29.69)*** | 1.981 (56.19)*** | -0.223 (-25.79)*** | 0.57 |

Panel B: Regression results when market declines ($R_{m,t} < 0$)

| <i>Market</i> | <i>C</i> | $ R_{m,t} $ | $R_{m,t}^2$ | \bar{R}^2 |
|---------------|---------------------|---------------------|-----------------------|-------------|
| Australia | 0.409 (36.30)*** | 1.814 (57.07)*** | -0.172 (-32.62)*** | 0.54 |
| China | 0.232 (21.57)*** | 1.065 (61.89)*** | -0.105 (-33.14)*** | 0.67 |
| Hong Kong | 0.388 (31.59)*** | 1.433 (67.80)*** | -0.128 (-37.49)*** | 0.65 |
| Indonesia | 0.283 (26.93)*** | 1.638 (78.56)*** | -0.170 (-38.17)*** | 0.75 |
| Japan | 0.281 (26.73)*** | 1.562 (64.55)*** | -0.236 (-36.34)*** | 0.64 |
| South Korea | 0.356 (25.73)*** | 1.390 (69.72)*** | -0.119 (-39.37)*** | 0.69 |
| Malaysia | 0.397 (36.47)*** | 1.066 (71.66)*** | -0.074 (-40.40)*** | 0.67 |
| Singapore | 0.318 (30.13)*** | 1.606 (75.66)*** | -0.196 (-39.71)*** | 0.73 |
| Thailand | 0.212 (26.72)*** | 1.588 (80.62)*** | -0.213 (-40.48)*** | 0.73 |
| Taiwan | 0.294 (25.47)*** | 1.455 (64.41)*** | -0.215 (-37.21)*** | 0.67 |
| United States | 0.449 (27.39)*** | 1.771 (56.32)*** | -0.201 (-30.64)*** | 0.58 |

Panel C: Wald test $H_0: \beta_{21} - \beta_{22} = 0$

| | $\beta_{21} - \beta_{22}$ | Chi-square statistics |
|---------------|---------------------------|-----------------------|
| Australia | -0.216 | (229.60)*** |
| China | -0.011 | (4.59)** |
| Hong Kong | -0.183 | (402.30)*** |
| Indonesia | -0.017 | (6.79)*** |
| Japan | 0.071 | (63.85)*** |
| South Korea | -0.040 | (53.14)*** |
| Malaysia | 0.012 | (26.87)*** |
| Singapore | 0.007 | (1.09) |
| Thailand | 0.005 | (0.48) |
| Taiwan | -0.032 | (13.74)*** |
| United States | -0.022 | (4.13)** |

Table 4. A summary of statistics for herding coefficients: 7/2/1997 – 3/23/2009

| | AU | CN | HK | ID | JP | KR | MA | SG | TH | TW | US |
|-----------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| Mean | -0.259 | -0.077 | -0.296 | -0.232 | -0.267 | -0.172 | -0.207 | -0.176 | -0.243 | -0.154 | 0.040 |
| Median | -0.240 | -0.070 | -0.259 | -0.213 | -0.266 | -0.171 | -0.192 | -0.177 | -0.240 | -0.155 | 0.046 |
| Maximum | -0.042 | 0.073 | -0.033 | 0.045 | -0.033 | 0.081 | 0.011 | 0.036 | -0.028 | 0.022 | 0.168 |
| Minimum | -0.574 | -0.267 | -0.759 | -0.624 | -0.533 | -0.372 | -0.434 | -0.411 | -0.571 | -0.384 | -0.078 |
| Std. Dev. | 0.127 | 0.055 | 0.168 | 0.134 | 0.099 | 0.074 | 0.095 | 0.083 | 0.093 | 0.071 | 0.047 |
| Skewness | -0.576 | -0.531 | -1.070 | -0.742 | -0.255 | -0.092 | -0.260 | -0.187 | -0.681 | -0.334 | -0.087 |
| Kurtosis | 2.494 | 3.550 | 3.3803 | 3.124 | 2.662 | 2.899 | 2.307 | 3.097 | 4.147 | 3.067 | 2.659 |
| T | 3059 | 3059 | 3059 | 3059 | 3059 | 3059 | 3059 | 3059 | 3059 | 3059 | 2894 |

Table 5. Correlation matrix of herding coefficients: 7/2/1997– 12/31/2008

This table reports the correlation of time-varying herding coefficients for each individual market. Numbers in parentheses are t-statistics.

| Markets | AU | CN | HK | ID | JP | KR | MA | SG | TH | TW | US |
|---------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|-------------------|-----------------|
| AU | 1.0000 ----- | | | | | | | | | | |
| CN | 0.3666 (21.789) | 1.0000 ----- | | | | | | | | | |
| HK | 0.4638 (28.948) | 0.0678 (3.7609) | 1.0000 ----- | | | | | | | | |
| ID | 0.2930 (34.499) | 0.2254 (12.792) | 0.4289 (26.256) | 1.0000 ----- | | | | | | | |
| JP | 0.4432 (27.341) | 0.4009 (24.198) | 0.1831 (10.298) | 0.4307 (26.389) | 1.0000 ----- | | | | | | |
| KR | 0.2410 (13.731) | 0.1000 (5.559) | 0.3238 (18.926) | 0.3549 (20.992) | 0.1221 (6.802) | 1.0000 ----- | | | | | |
| MA | 0.1971 (11.121) | -0.0720 (-3.992) | 0.5691 (38.266) | 0.5955 (40.985) | 0.0832 (4.617) | 0.3174 (18.508) | 1.0000 ----- | | | | |
| SG | 0.3266 (19.110) | 0.1058 (5.883) | 0.6744 (50.502) | 0.3192 (18.624) | 0.1754 (9.852) | 0.5300 (34.559) | 0.4946 (31.467) | 1.0000 ----- | | | |
| TH | 0.0788 (4.373) | 0.0986 (5.478) | 0.2887 (16.675) | -0.0169 (-0.936) | 0.1206 (6.717) | 0.1984 (11.196) | -0.0108 (-0.596) | 0.3331 (19.532) | 1.0000 ----- | | |
| TW | 0.0081 (0.449) | 0.2039 (11.517) | 0.1042 (5.796) | -0.0267 (-1.482) | 0.3059 (17.764) | 0.2441 (13.92) | 0.0270 (1.498) | 0.2456 (14.01) | 0.2808 (16.17) | 1.0000 ----- | |
| US | -0.4505 (-27.90) | -0.0475 (-2.631) | -0.4591 (-28.57) | -0.5582 (-37.20) | -0.1678 (-9.413) | -0.3358 (-19.71) | -0.4286 (-26.23) | -0.2753 (-15.83) | -0.1660 (-9.311) | 0.0371 (2.054) | 1.0000 ----- |

Table 6. Factors that determine dynamic herding behavior for sample period: 1/4/1999 – 3/23/2009

This table reports consistent estimates for the herding measure based on the following regression model:

$$Herd_t = \phi_0 + \phi_1 R_{m,t} + \phi_2 \hat{\sigma}_t^2 + \varepsilon_t$$

where $Herd_t$ is the estimated herding coefficient, $\beta_{2,t}$, derived from the state-space model; $R_{m,t}$ is stock market return, $\hat{\sigma}_t^2$ is the conditional variance of stock returns based on an asymmetric GARCH(1,1) process for each market, ϕ_i ($i = 0, 1, 2$) denotes constant parameters. *** indicates significance at the 1% level.

| Markets | Hypothesis 1 | | | Hypothesis 2 | | | Joint Hypothesis | | | |
|-----------|------------------------|----------------------|-------------|------------------------|---------------------|-------------|------------------------|----------------------|---------------------|-------------|
| | C | $R_{m,t}$ | \bar{R}^2 | C | $\hat{\sigma}_t^2$ | \bar{R}^2 | C | $R_{m,t}$ | $\hat{\sigma}_t^2$ | \bar{R}^2 |
| Australia | -0.273 (-109.26)*** | -0.029 (-7.40)*** | 0.02 | -0.278 (-96.18)*** | 0.020 (4.10)*** | 0.03 | -0.319 (-65.20)*** | -0.022 (-6.38)*** | 0.012 (2.86)*** | 0.08 |
| China | -0.077 (-72.09)*** | -0.004 (-7.28)*** | 0.01 | -0.098 (-59.79)*** | 0.006 (21.13)*** | 0.11 | -0.104 (-61.93)*** | -0.004 (-8.53)*** | 0.004 (13.20)*** | 0.17 |
| HK | -0.319 (-97.92)*** | -0.008 (-2.76)*** | 0.01 | -0.357 (-85.45)*** | 0.029 (13.68)*** | 0.15 | -0.375 (-51.61)*** | -0.011 (-4.03)*** | 0.031 (19.48)*** | 0.16 |
| Indonesia | -0.232 (-95.67)*** | -0.005 (-2.36)** | 0.01 | -0.272 (-83.88)*** | 0.023 (14.95)*** | 0.10 | -0.282 (-66.33)*** | -0.008 (-3.42)*** | 0.026 (19.33)*** | 0.13 |
| Japan | -0.276 (-143.24)*** | -0.006 (-2.79)*** | 0.01 | -0.293 (-106.49)*** | 0.017 (8.46)*** | 0.07 | -0.337 (-97.92)*** | -0.005 (-2.96)*** | 0.012 (6.63)*** | 0.14 |
| S. Korea | -0.174 (-118.88)*** | -0.006 (-5.76)*** | 0.01 | -0.193 (-112.75)*** | 0.008 (16.52)*** | 0.13 | -0.200 (-93.84)*** | -0.006 (-6.13)*** | 0.007 (14.73)*** | 0.17 |
| Malaysia | -0.218 (-117.48)*** | 0.000 (0.13) | 0.01 | -0.242 (-101.74)*** | 0.012 (12.63)*** | 0.19 | -0.218 (-66.61)*** | -0.001 (-0.79) | 0.011 (13.19)*** | 0.24 |
| Singapore | -0.180 (-107.67)*** | 0.000 (0.04) | 0.01 | -0.212 (-102.08)*** | 0.024 (22.87)*** | 0.22 | -0.221 (-69.77)*** | 0.000 (0.18) | 0.024 (24.02)*** | 0.23 |
| Thailand | -0.252 (-139.36)*** | -0.010 (-3.93)*** | 0.01 | -0.265 (-119.58)*** | 0.020 (10.43)*** | 0.05 | -0.282 (-113.20)*** | -0.009 (-4.04)*** | 0.011 (6.34)*** | 0.12 |
| Taiwan | -0.154 (-112.50)*** | -0.007 (-7.40)*** | 0.02 | -0.169 (-86.92)*** | 0.007 (10.38)*** | 0.03 | -0.166 (-43.10)*** | -0.006 (-6.60)*** | 0.004 (5.43)*** | 0.11 |

Table 7. Factors that determine dynamic herding behavior: Stock returns, conditional volatility, cross-market herding, and lagged VIX (sample period: 1/4/1999 – 3/23/2009)

This table reports regression estimates of herding dynamics by using current stock returns ($R_{m,t}$), conditional variance ($\hat{\sigma}_t^2$), cross-market herding, and lagged global market volatility (VIX_{t-1}). The estimated equation is given by:

$$Herd_t = \phi_0 + \phi_1 R_{m,t} + \phi_2 \hat{\sigma}_t^2 + \phi_3 (Cross\ HERD)_t + \phi_4 VIX_{t-1} + \varepsilon_t$$

***, **, and * indicate significance at the 1%, 5%, and 10% levels, respectively.

| Markets | Domestic and cross-market independent variables | | | | | |
|-------------|---|----------------------|---------------------|---------------------|-----------------------|-------------|
| | C | $R_{m,t}$ | $\hat{\sigma}_t^2$ | $Cross\ HERD_t$ | VIX_{t-1} | \bar{R}^2 |
| Australia | -0.198 (-22.18)*** | -0.018 (-6.00)*** | 0.008 (2.70)*** | 0.332 (26.41)*** | -0.001 (-4.50)*** | 0.24 |
| China | -0.045 (-11.91)*** | -0.004 (-9.21)*** | 0.005 (15.62)*** | 0.210 (19.45)*** | 0.001 (2.07)** | 0.29 |
| Hong Kong | -0.193 (-18.21)*** | -0.007 (-3.06)*** | 0.017 (13.50)*** | 1.145 (39.18)*** | 0.002 (6.39)*** | 0.54 |
| Indonesia | -0.133 (-17.50)*** | -0.004 (-2.75)*** | 0.006 (6.27)*** | 0.797 (30.37)*** | 0.001 (6.83)*** | 0.38 |
| Japan | -0.223 (-32.86)*** | -0.005 (-3.08)*** | 0.018 (14.04)*** | 0.613 (19.25)*** | -0.002 (-11.43)*** | 0.28 |
| South Korea | -0.247 (-51.49)*** | -0.003 (-4.23)*** | 0.002 (0.89) | 0.060 (6.46)*** | 0.004 (24.30)*** | 0.26 |
| Malaysia | -0.130 (-20.35)*** | -0.001 (-0.65) | 0.006 (9.74)*** | 0.462 (25.13)*** | 0.001 (2.39)** | 0.40 |
| Singapore | -0.141 (-22.53)*** | 0.002 (1.71)* | 0.011 (18.00)*** | 0.276 (34.00)*** | 0.001 (7.74)*** | 0.54 |
| Thailand | -0.233 (-37.16)*** | -0.009 (-4.10)*** | 0.004 (2.35)** | 0.306 (18.27)*** | 0.000 (2.29)** | 0.20 |
| Taiwan | -0.129 (-25.69)*** | -0.006 (-6.66)*** | 0.004 (4.29)*** | 0.176 (13.42)*** | 0.001 (2.14)** | 0.16 |

Table 8. Correlation analysis of VIX and conditional variance

This table reports the correlation statistics between $\hat{\sigma}_t^2$ (conditional variance) and VIX_{t-1} . The first column shows the correlation coefficients ranging from 0.16 to 0.50, with the highest value in South Korea. The numbers in parentheses are t-statistics.

| t-Statistic | VIX_{t-1} | $\hat{\sigma}_t^2$ (AU) | $\hat{\sigma}_t^2$ (CN) | $\hat{\sigma}_t^2$ (HK) | $\hat{\sigma}_t^2$ (ID) | $\hat{\sigma}_t^2$ (JP) | $\hat{\sigma}_t^2$ (KR) | $\hat{\sigma}_t^2$ (MA) | $\hat{\sigma}_t^2$ (SG) | $\hat{\sigma}_t^2$ (TH) | $\hat{\sigma}_t^2$ (TW) |
|-------------------------|--------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|-------------------------|
| VIX_{t-1} | 1.0000 ----- | | | | | | | | | | |
| $\hat{\sigma}_t^2$ (AU) | 0.2508 (13.375) | 1.0000 ----- | | | | | | | | | |
| $\hat{\sigma}_t^2$ (CN) | 0.2778 (14.927) | 0.1352 (7.045) | 1.0000 ----- | | | | | | | | |
| $\hat{\sigma}_t^2$ (HK) | 0.3933 (22.082) | 0.2787 (14.979) | 0.2486 (13.249) | 1.0000 ----- | | | | | | | |
| $\hat{\sigma}_t^2$ (ID) | 0.1837 (9.647) | 0.0861 (4.465) | 0.0786 (4.072) | 0.3324 (18.194) | 1.0000 ----- | | | | | | |
| $\hat{\sigma}_t^2$ (JP) | 0.4298 (24.573) | 0.2215 (11.728) | 0.1846 (9.699) | 0.5205 (31.468) | 0.1648 (8.625) | 1.0000 ----- | | | | | |
| $\hat{\sigma}_t^2$ (KR) | 0.5015 (29.921) | 0.2651 (14.196) | 0.0451 (2.335) | 0.5175 (31.221) | 0.3546 (19.580) | 0.3828 (21.393) | 1.0000 ----- | | | | |
| $\hat{\sigma}_t^2$ (MA) | 0.1615 (8.450) | 0.0608 (3.1439) | 0.0201 (1.040) | 0.3447 (18.955) | 0.4223 (24.051) | 0.1172 (6.092) | 0.5004 (29.836) | 1.0000 ----- | | | |
| $\hat{\sigma}_t^2$ (SG) | 0.2946 (15.913) | 0.1501 (7.840) | 0.2131 (11.258) | 0.4855 (28.669) | 0.5207 (31.483) | 0.3091 (16.776) | 0.4865 (28.742) | 0.5793 (36.688) | 1.0000 ----- | | |
| $\hat{\sigma}_t^2$ (TH) | 0.2905 (15.670) | 0.1160 (6.031) | 0.0789 (4.088) | 0.3872 (21.676) | 0.3228 (17.605) | 0.4652 (27.126) | 0.4759 (27.932) | 0.3915 (21.964) | 0.3748 (20.871) | 1.0000 ----- | |
| $\hat{\sigma}_t^2$ (TW) | 0.4716 (27.607) | 0.2016 (10.624) | 0.0727 (3.764) | 0.4011 (22.602) | 0.1210 (6.295) | 0.2733 (14.665) | 0.4517 (26.135) | 0.2444 (13.013) | 0.2973 (16.075) | 0.1975 (10.403) | 1.0000 ----- |

Table 9. Estimates of herding equation with different specifications of the nonlinear components

This table reports the regression results of CSAD. The estimated equation is:

$$CSAD_t = b_0 + b_1|R_{m,t}| + b_2R_{m,t}^2 + b_3R_{m,t}^3 + b_4\hat{\sigma}_t^2R_{m,t}^2 + \varepsilon_t \quad (8)$$

where $CSAD_t$ is equally weighted cross-sectional absolute deviations; $R_{m,t}$ is the value of an equally weighted realized return of all industry indexes on day t ; and $\hat{\sigma}_t^2$ is the conditional variance of $R_{m,t}$. The data in this study are daily observations from individual firms for each market, spanning from 7/2/1997 to 3/23/2009. \bar{R}^2 is the adjusted R -squared. The numbers in parentheses are t-statistics. F statistic is used to test the restriction: $b_3 = b_4 = 0$. The rightmost column is χ^2 test for $b_2 - \beta_2 = 0$, where β_2 is the coefficient t from the model

$$CSAD_t = \beta_0 + \beta_1|R_{m,t}| + \beta_2R_{m,t}^2 + \varepsilon_t \quad (2)$$

***, **, * denote that the coefficient is significant at the 1%, 5%, and 10% level, respectively.

| Markets | Explanatory Variables | | | | | | | |
|-------------|-----------------------|---------------------|-----------------------|----------------------|------------------------------------|-------------|-----------------------------|---------------------------------|
| | C | $ R_{m,t} $ | $R_{m,t}^2$ | $R_{m,t}^3$ | $R_{m,t}^2 \cdot \hat{\sigma}_t^2$ | \bar{R}^2 | F- Test for $b_3 = b_4 = 0$ | χ^2 $b_2 - \beta_2 = 0$ |
| Australia | 1.208 (71.63)*** | 0.918 (21.78)*** | -0.119 (-6.24)*** | -0.005 (-3.37)*** | -0.002 (-0.46) | 0.49 | (7.21)*** | -0.059 (20.25)*** |
| China | 0.927 (26.74)*** | 0.461 (14.67)*** | -0.048 (-7.13)*** | -0.001 (-4.35)*** | 0.002 (3.63)*** | 0.36 | (14.38)*** | -0.018 (10.56)*** |
| Hong Kong | 1.228 (53.34)*** | 0.931 (26.56)*** | -0.129 (-11.79)*** | -0.004 (-4.67)*** | 0.002 (4.83)*** | 0.54 | (17.29)*** | -0.071 (39.53)*** |
| Indonesia | 0.881 (70.90)*** | 1.074 (51.04)*** | -0.101 (-17.42)*** | 0.001 (2.70)** | 0.001 (2.23)** | 0.68 | (10.75)*** | -0.017 (4.75)** |
| Japan | 1.060 (44.30)*** | 0.614 (21.90)*** | -0.075 (-8.540)*** | 0.001 (1.79)* | 0.002 (3.12)*** | 0.35 | (13.46)*** | -0.018 (8.56)*** |
| South Korea | 1.411 (54.44)*** | 0.535 (22.84)*** | -0.023 (-4.48)*** | 0.001 (2.47)** | 0.000 (0.56) | 0.54 | (6.69)*** | 0.038 (1.80) |
| Malaysia | 1.011 (76.72)*** | 0.746 (43.13)*** | -0.073 (-17.20)*** | 0.000 (-1.80)* | 0.001 (5.41)*** | 0.67 | (16.81)*** | -0.042 (12.98)*** |
| Singapore | 1.091 (64.84)*** | 0.764 (32.64)*** | -0.079 (-11.86)*** | -0.001 (-1.86)* | 0.001 (2.52)** | 0.64 | (3.35)** | -0.020 (3.19)* |
| Thailand | 0.667 (87.21)*** | 0.939 (55.11)*** | -0.125 (-18.26)*** | -0.004 (-6.52)*** | 0.001 (1.92)** | 0.67 | (22.80)*** | -0.028 (18.52)*** |
| Taiwan | 1.140 (76.06)*** | 0.689 (33.36)*** | -0.111 (-16.06)*** | -0.001 (-2.79)** | 0.005 (4.68)*** | 0.40 | (13.56)*** | -0.027 (9.67)*** |

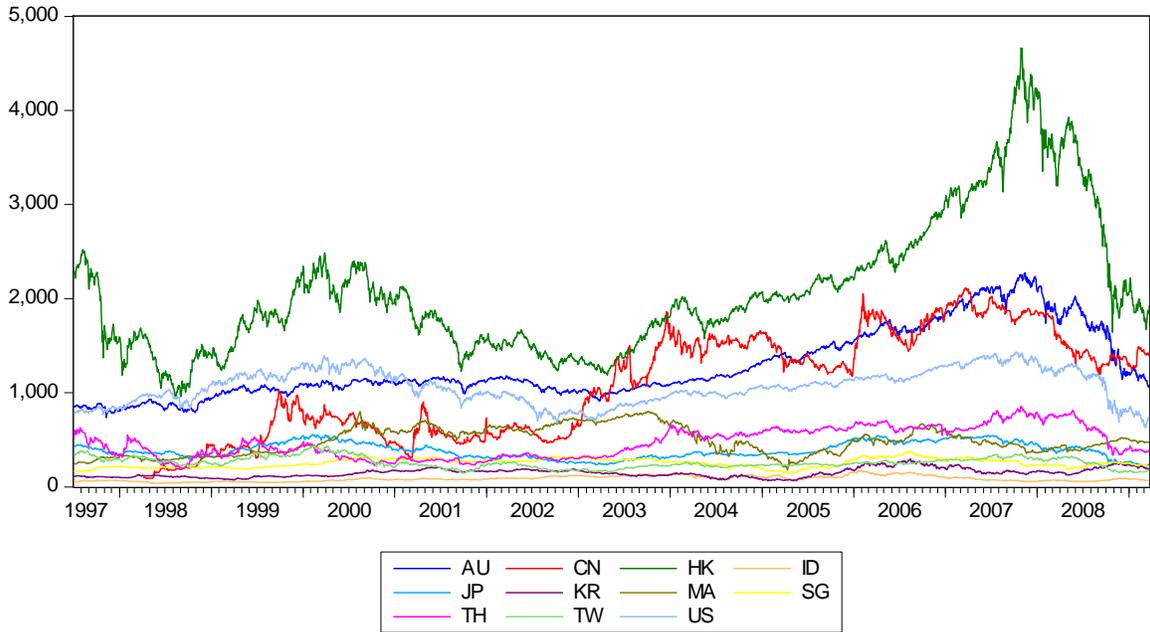


Figure 1. Time series plots of aggregate market index based on Datastream data

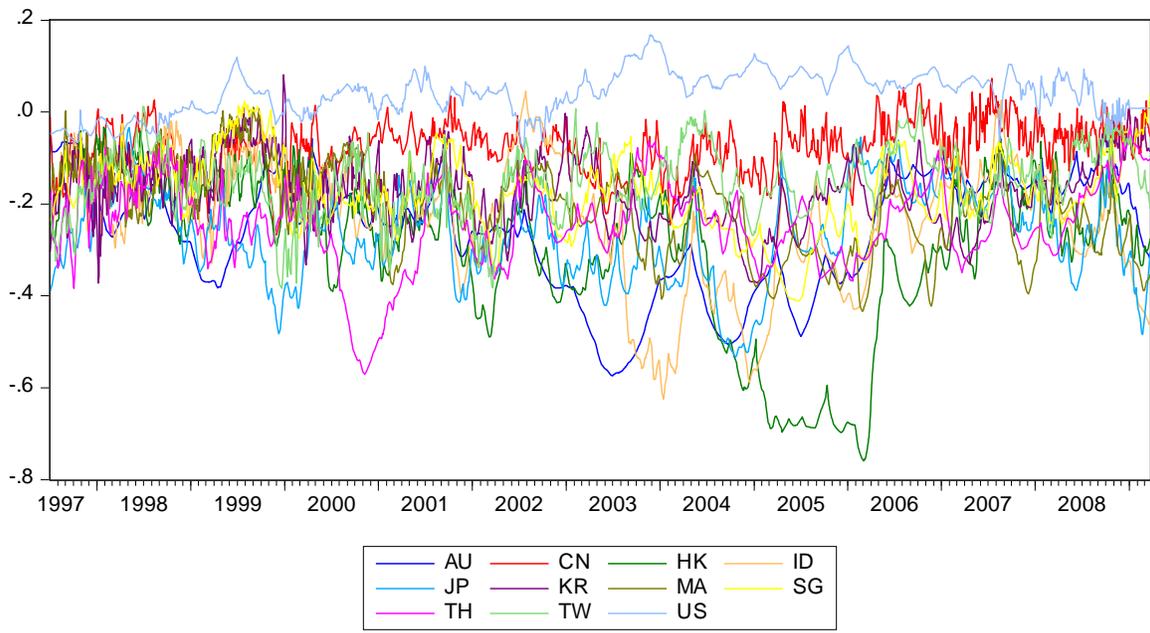


Figure 2. Time series plots of herding coefficients for 11 markets